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The Minimum Wage and Youth Unemployment

By Steven P. Zell

Of all of the factors involved in the problem of youth unemployment, none has been more extensively debated than the effects of the minimum wage. Governed by the Fair Labor Standards Act of 1938 and periodic amendments to it, the level of the minimum wage and its coverage are once again an important issue. The 1977 amendments raised the minimum wage from the previous \$2.30 per hour to \$2.65, effective January 1, 1978, and future scheduled increases will raise the level to \$3.35 by 1981. Table 1 presents a history of the statutory changes in the minimum wage from 1938 to date.

The changes in the minimum wage are receiving special attention at this time because of the high level of unemployment in the economy, especially for minority teenagers. In recent years, economists have conducted scores of econometric studies to try to ascertain the effects of changes in the minimum wage on the labor market. For the most part, the studies have focused on teenagers, with most of the emphasis on the unemployment impact. Some other studies, however, have examined the effect on employment, and at least two have looked at the impact of the minimum

wage on the distribution of income.

Given all of this research, a consensus might reasonably have been expected to emerge that a given percentage change in the minimum wage would result, on average, in a certain percentage change in unemployment or employment. Unfortunately, this is not the case. Although the majority of the sophisticated models do find significant disemployment effects for the minimum wage (that is, a reduction in employment in response to an increase in the minimum wage), other valid approaches yield **inconclusive** results. Furthermore, of those studies yielding the theoretically expected negative impact of the minimum wage, the range of estimated effects is sufficiently large as to make the estimates of uncertain policy value.

This article analyzes the reasons for this lack of consensus. The article begins with the development of two alternative models of the labor market, which are used to analyze the theoretical impact of the minimum wage on employment and unemployment. The theoretical approach is then contrasted with the problems involved in the empirical estimation of the minimum wage impact. Finally, in light of these findings, the article concludes with

a discussion of some of the important new evidence about minimum wage issues?

WHAT'S WRONG WITH THE MINIMUM WAGE (AND WHY IS IT SO HARD TO PROVE?)

When viewed abstractly, the minimum wage might appear to be a reasonably good tool for protecting poor workers from unconscionably low wages, helping to relieve poverty, and maintaining purchasing power in the economy? Economic analysis reveals, however, that because of a number of unintended effects of the minimum wage, these benefits do not necessarily result.

The Theoretical Model

The theoretical impact of the minimum wage is fairly simple and straightforward. If there were no imperfections in the labor market, if wages could move both up and down, and if all adjustments were instantaneous, the situation would be approximated by the supply curve, SS, and the demand curve, DD, as shown in Figure 1. The supply curve shows the number of workers who would be willing to work at any given wage, the demand curve shows the number who would be demanded by employers at any wage, and their intersection represents an **equilib-**

¹ The best evaluation of minimum wage research through 1974 is Robert S. Goldfarb, "The Policy Content of Quantitative Minimum Wage Research," **Proceedings of the 27th Annual Meeting of the Industrial Relations Research Association, 1974**, pp. 261-68. See also **Minimum Wage and Maximum Hour Standards Under the Fair Labor Standards Act**, Employment Standards Administration, U.S. Department of Labor, October 1977, pp. 41-54; and **The Problem of Teenage Unemployment**, Alan A. Fisher, unpublished Ph.D. dissertation, University of California, Berkeley, August 1973, Chapter 3.

² See Clara F. Schloss, "Closing the Minimum Wage Gap," **AFL-CIO American Federationist**, January 1976.

**Table 1
MINIMUM WAGE RATE CHANGES**

	Nonfarm Workers		Farm Workers*
	Covered	Newly Covered	
October 24, 1938	\$0.25	\$ -	\$ -
October 24, 1939	0.30	-	-
October 24, 1945	0.40	-	-
January 25, 1950	0.75	-	-
March 1, 1956	1.00	-	-
September 3, 1961	1.15	1.00	-
September 3, 1963	1.25	-	-
September 3, 1964	-	1.15	-
September 3, 1965	-	1.25 [†]	-
February 1, 1967	1.40 [†]	1.00	1.00
February 1, 1968	1.60	1.15	1.15
February 1, 1969	-	1.30	1.30
February 1, 1970	-	1.45	-
February 1, 1971	-	1.60	-
May 1, 1974	2.00	1.90	1.60
January 1, 1975	2.10	2.00	1.80
January 1, 1976	2.30	2.20	2.00
January 1, 1977	-	2.30	2.20
January 1, 1978	2.65	2.65	2.65
January 1, 1979	2.90	2.90	2.90
January 1, 1980	3.10	3.10	3.10
January 1, 1981	3.35	3.35	3.35

*Not all farm workers are covered by the minimum wage.

[†]All job categories covered prior to the 1966 amendments were raised to \$1.40 per hour in February 1967.

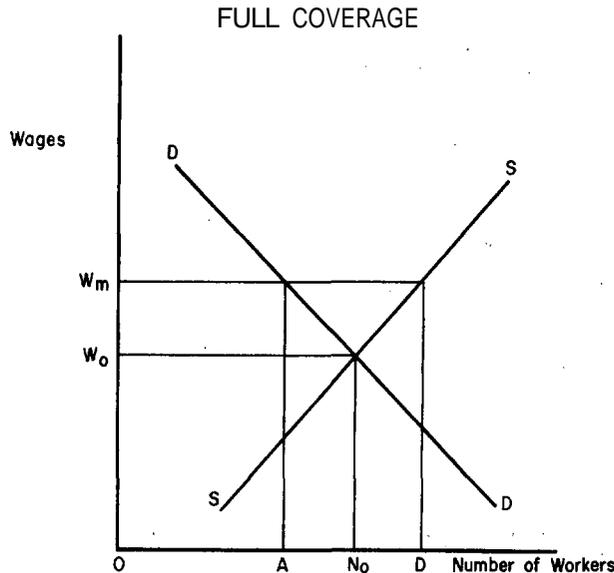
SOURCE: 1938-1971: Bureau of Labor Statistics, **Youth Unemployment and Minimum Wages**, Bulletin 1657, Washington, D.C., GPO, 1970, p. 11, Table 1.5, **1974-1977:** U.S. Department of Labor, Employment Standards Administration, **Minimum Wage and Maximum Hour Standards Under the Fair Labor Standards Act**, an Economic Effects Study submitted to Congress 1977, p. 3.

rium where the number of workers actually hired, given by distance ON_0 , would receive the market wage, OW_0 .

Suppose now that a minimum wage, OW_m , which is greater than the equilibrium wage, OW_0 , is imposed on the labor market. Since at the higher minimum wage, the number of workers

Figure 1

One Sector Labor Market



who wish to work, OD , exceeds the number of workers that will be demanded by business, OA , some workers, AD , will be without jobs. If all of these unemployed workers stay in the labor market, the unemployment rate—the number of unemployed workers divided by the number in the labor market—will be AD/OD , an involuntary unemployment caused by the imposition of the minimum wage.

A somewhat more realistic picture is presented in Figures 2a and 2b by relaxing some of the previous assumptions. Here assume that the economy consists of two sectors: a covered sector, in which workers receive the minimum wage (Figure 2a); and an uncovered sector, in which wages are set by supply and demand (Figure 2b). Assume also that people have the option of working (or seeking work) in either sector, that those not working may be either unemployed or out of the labor force, and that

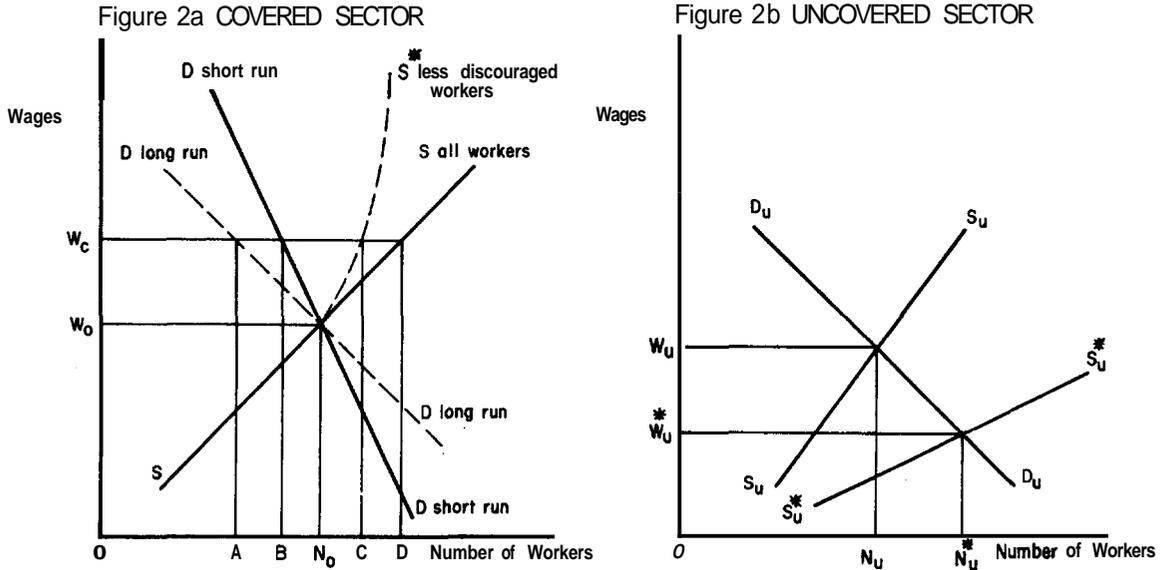
adjustments to outside influences (as the imposition or increase of a minimum wage) take time?

Initially, in the absence of a minimum wage, there is no unemployment (both markets clear), and those employed in the covered market, ON_0 , receive the covered market wage, OW_0 . When a minimum wage is imposed in the covered sector, demand first reacts along D *short run*. Some firms cut back on their labor immediately, but others adjust only slowly. However, because the demand curve has a negative slope, fewer workers, only OB , are hired at the minimum wage, leaving a number, BD , involuntarily unemployed, and raising the unemployment rate, from zero to BD/OD . Over time, the adjustment proceeds further. Some unskilled labor is replaced by higher skilled labor and capital, and the relevant demand curve is D *long run*.

³ Figure 2a is taken from Fisher, p. 53.

Figure 2

Two Sector Labor Market



If supply were not affected, unemployment would rise, to AD, the unemployment rate would also climb, to ADIOD, and employment would fall further, to OA.

However, supply effects are also likely, and these can work in two directions. The higher posted wage in the covered sector could lure some workers from the uncovered to the covered sector. On the other hand, because the probability of finding employment has fallen, workers will tend to leave the covered sector and either drop out of the labor force or enter the uncovered sector. The most likely net effect is a fall in supply, shown by a leftward shift in the supply curve to S^* , yielding a lower level of unemployment, AC, and a lower unemployment rate, ACIOC, than before the supply shift. Covered employment, OA, is not further changed by the supply shift, and remains below the equilibrium level. However, the

imposition of the minimum wage adversely affects an additional number of workers, CD, by forcing them out of the higher wage covered sector.

The uncovered sector is also affected by the minimum wage, though indirectly. Generally viewed as consisting of small, low-wage firms, the uncovered sector may be pictured as in Figure 2b, with an equilibrium wage, OW_U , lower than the covered equilibrium wage, OW_0 . If the net effect of the minimum wage is a shift of workers into the uncovered sector, this results in an increase in the supply of workers or a rightward shift in the uncovered supply curve to S^*_U . Since the new uncovered equilibrium is given by the intersection of the demand curve with the new supply curve, total uncovered employment increases, but the new wage rate is lower for all uncovered workers. In general, the final effect of the minimum wage imposition (or increase) will be

lower covered employment, greater covered unemployment, higher uncovered employment (though not enough to absorb all the covered loss), and a lower uncovered wage rate?

Problems of Estimation

Given this scenario, why is it so difficult to estimate the actual effects of the minimum wage? Primarily, the difficulty arises because the labor market is far more complex than can be exactly modeled by a pair of diagrams or even the most sophisticated set of equations. In particular, an analysis of the type in diagrams 2a and 2b implicitly assumes that the labor market can be readily divided into a single covered and a single uncovered sector, each of which contains a homogeneous type of labor receiving its respective wage rate. The real world, of course, is much more complicated, consisting of many different types of labor and many different equilibrium wage rates. Nevertheless, in order to estimate the impact of the imposition or increase in the minimum wage, a researcher must separate and model that part of the labor market in which workers are essentially homogeneous and in which the statutory minimum wage, W_C in Figure 2a, exceeds W_O , the market clearing wage, thereby setting an effective floor on wages once it is imposed. Rather than being an identifiable part of the economy,

⁴ Using a general equilibrium analysis, James F. Ragan, Jr., points out that, theoretically, the change in relative wage costs in the two sectors can result in a shift in product and derived labor demand so as to result in a net increase in total employment. Among other things, this result assumes that most of the discouraged covered workers would be willing to shift to the uncovered sector. See Ragan, "The Theoretical Ambiguity of a Minimum Wage," *Atlantic Economic Journal*, March 1977.

however, such a sector consists of parts of some industries and occupations, and includes some teenagers (but not all) and some adults as well. In fact, although a greater percentage of teenagers than adults are low-wage recipients, adults are a much larger group, and by far the greatest number of those receiving low wages are **adults**.⁵

The difficulty in identifying the appropriate covered sector highlights the fact that, in an economically meaningful sense, there is probably no such thing as a homogeneous "teenage labor market." From a theoretical standpoint, it would be preferable to consider the covered sector as consisting of a low productivity labor market in which the minimum wage sets an effective wage floor (and where, incidentally, teenagers are concentrated) and a high productivity, high-wage labor market?

All studies which, nevertheless, attempt to estimate the impact of a minimum wage change on "the teenage labor market," implicitly assume that an economically meaningful "teenage labor market" actually exists. The market so estimated, however, is both too narrow and too broad: too narrow in that the more theoretically appropriate low-productivity labor market includes many nonteenagers who also compete for the low-wage jobs, and too broad because many of the teenagers included, especially those 18-19 years of age, are actually able to find high-wage

⁵ See Edward M. Gramlich, "The Impact of the Minimum Wage on Other Wages, Employment, and Family Incomes," *Brookings Papers on Economic Activity*, 1978: 2, Tables 2 and 3, and Fisher, pp. 116-18 and Tables 2.7-2.9.

⁶ My thanks to James F. Ragan, Jr., for his helpful comments on this point. The low productivity sector would correspond to diagram 2a. The high productivity sector has no counterpart in the two diagrams.

employment and thus are not directly affected by the minimum wage?

However, though theoretically superior, the concept of a low-productivity labor market is almost impossible to work with. Given available data, there is simply no way to identify those persons who should be included in the low-wage population over which the model is to be estimated. On empirical grounds, then, the researcher is largely forced to work with some measure of a teenage labor market. In so doing, however, he commits potentially serious econometric errors. These errors arise because the data used in the estimation procedure are not strictly appropriate to explain the minimum wage impact on the employment of teenagers. For example, if a teenage wage variable were available and used as one of the explanatory variables, it would be misspecified because it would be an average of wages received by both high- and low-wage teenagers. Furthermore, since teenage employment depends upon the wage rate in both high- and low-wage sectors, including only one wage variable means that a variable has been omitted from the model. Both of these specification errors tend to bias the estimated results. A similar problem arises from not considering the impact on teenage employment of the wage distribution of low-wage adults who compete with low-wage teenagers for low-productivity jobs⁸

In general, however, researchers have been forced to ignore these problems. Instead, the format that most models finally take is to try to explain, through regression analysis, some variant of teenage employment or unemployment (the dependent variable) by several labor market factors such as the minimum

wage, a measure of economic activity, and some other independent variables. Using time series data, the model is generally of the form:

$$Z_t = f(AD, MW, X_1, \dots, X_n)$$

Z_t is either the teenage unemployment rate or the ratio of teenage unemployment or employment to the teenage population for a specific age-sex-color group.

AD is a measure of aggregate demand, like the gap between real and potential GNP, or a measure of labor market tightness, like the adult male unemployment rate.

MW is a measure of the minimum wage.

X_1, \dots, X_n are other independent variables such as a time trend, a coverage variable for the minimum wage, or the supply of teenagers.

f indicates the functional form, generally linear or log-linear.

The fundamental difficulty involved in specifying this model for estimation, however, is that the available data are of extremely poor quality. Economic theory has far surpassed the state of the data. As a result, extremely sophisticated models, including all of the "right" variables can be elaborated but cannot be estimated as originally presented. Instead, serious compromises must be made in which some desired variables are omitted, proxies are used for other variables for which data do not exist, or

⁷ Fisher, p. 115.

⁸ Fisher, pp. 115-16.

data series are generated through extremely complicated, but often questionable, processes.

The Case of the Minimum Wage Variable

The most striking example of this data problem concerns the choice of MW, the minimum wage variable. If an estimate is to be made, using time series data, of the effect of the minimum wage on teenage employment, the minimum wage data series must first have the influence of inflation removed from it. Clearly, a \$2.30 minimum wage would affect employment very differently in 1938 than in 1977. The correct deflation factor, however, is probably not a series like the consumer price index. Instead, according to the earlier theoretical discussion, the proper choice is the market clearing wage for teenagers (W_0 in Figure 2a). The reason is that, given the covered labor market's supply and demand curves, the level of unemployment induced by the minimum wage is determined by the relationship between the minimum wage and the market clearing wage that would have existed in its absence. Unfortunately, the market clearing wage series for teenagers is unknown. Other price series may be used as the deflator in its place only if their movements over time tend to parallel those of the market wage.

The method most researchers use to deflate their minimum wage variables is to divide the statutory minimum wage, MW^* , by the average hourly earnings of nonsupervisory employees in either manufacturing or the private nonfarm economy (AHE), thus yielding the variable MW in the model described earlier. The principal attractiveness of AHE is that it is a reasonably accurate, readily available time series estimate of wages. But it is not likely that AHE moves, over time, like the hypothetical and unknown market-clearing wage for

teenagers. Nor is it likely that AHE is an equally good approximation for the various market wages that exist in the many industries or occupations that employ teenagers. Because its denominator has no necessary relationship to the teenage labor market, the **inflation-adjusted** minimum wage variable will tend to have attributed to it a biased estimate of its influence on the employment of teenagers.

The minimum wage variable used by many studies has a further complication. Since the minimum wage was first introduced in 1938, two types of changes have been taking place: hourly wages have been rising, and the size of the uncovered sector has been falling as new industries are covered, and industries previously covered in part experience expanded coverage. Furthermore, when workers become newly covered, they generally enter coverage at a lower minimum wage, and through larger or more frequent increases, gradually achieve the regular schedule (Table 1).

Because changes in coverage clearly affect the impact of the minimum wage, these changes must be taken into account in the model. The way this has generally been done is to define a new, weighted minimum wage variable, which may be called **MWAGE**.⁹

If there were only one industry and all of the jobs in it had been covered by the minimum wage for many years, the variable MWAGE would correspond to the variable MW discussed earlier. However, this is not the case. The variable MWAGE, instead, incorporates the fact that the total effect of the minimum wage on employment is the weighted sum of the different effects on employment by industry. The weights used recognize that not only do some industries have partial coverage, but that industries also vary as to the percentage of workers in job

categories that have been covered for many years and the percentage newly covered by recent amendments. Because the **MWAGE** variable incorporates this weighting scheme as well as the adjustment for inflation, it is referred to as the "effective minimum wage." Though representing a major conceptual improvement over earlier methodologies that excluded a coverage variable, this composite variable also has serious problems that greatly complicate the interpretation and accuracy of the estimated equation.

9

$$MWAGE = \sum_i \left[\frac{E_i}{E_t} \left\{ \frac{MB_i}{AHE_i} \cdot CB_i \right\} + \frac{MN_i}{AHE_i} \cdot CN_i \right]$$

where **i** = major industry division.
t = total nonfarm economy.
E = payroll employment.
AHE = average hourly earnings of non-supervisory employees.
MB = basic minimum wage for long-time covered workers.
MN = minimum for newly covered workers.
CB = proportion of nonsupervisory employees covered by the **basic** minimum for long-time covered workers.
CN = proportion of nonsupervisory workers covered by the most recent **minimum** wage amendment.

This method was developed by the U.S. Bureau of Labor Statistics, **Youth Unemployment and Minimum Wages**, Bulletin 1657, Washington, D.C., Government Printing Office, 1970, p. 12. Essentially, what this variable does is to define an effective minimum wage as a double weighted average (MWAGE). Within each industry, denoted by **i**, the respective **minimum** wage rates of the long-time covered (**MB**) and newly covered (**MN**) workers are deflated by the level of average hourly earnings (**AHE**), and weighted by the proportion of nonsupervisory workers either long-time (**CB**) or newly covered (**CN**) in that industry. These two products are summed, and weighted in turn by the proportion of all payroll employment represented by that industry (E_i/E_t).

The difficulties engendered by the combined minimum **wage/coverage** variable (**MWAGE**) are basically of two kinds. First, the combined variable implicitly assumes that a given percentage change in coverage has exactly the same effect on the dependent variable as does the same percentage change in the wage level. There is, however, no theoretical basis for believing this to be the **case**.¹⁰ The second type of problem with this combined variable involves the weighting procedure used to apportion coverage and minimum wage changes among industries. Like the problems involved in the proper choice of a deflated wage variable—problems that still remain here—similar complications also arise from the use of coverage weights in **MWAGE** which are totally unrelated to the hypothetical teenage labor market. The important point is that, in order to explain changes in the dependent variable—for example, teenage employment—movements in independent variables like **MWAGE** must lead to (theoretically, should cause) movements in the dependent variable. If changes in minimum wage coverage are to affect teenage employment, therefore, the coverage changes included in the **MWAGE** variable must either be changes in the coverage of teenagers or, if this is

¹⁰ Furthermore, in a model developed to estimate the responsiveness of employment to minimum wage changes (the wage elasticity of employment), Gramlich has shown that the coverage variable and the basic minimum wage term should enter the equation with different coefficients and in different mathematical forms. Ignoring this fact, he concludes, probably biases downward the total estimated elasticity for the minimum wage and coverage changes. Gramlich, pp. 415 and 433. On the other hand, Michael C. Lovell, in "The Minimum Wage Reconsidered," **Western Economic Journal**, December 1973, reports that his particular model was not especially sensitive to his minimum wage variable's specification.

unknown, proxies for these changes. But the distribution of teenagers *among* industries, and the occupational distribution of teenagers *within* industries, differs greatly from that for "all nonsupervisory employees," the group to which the coverage weights (CB and CN in footnote 9) in MWAGE actually apply. Because of this, the coverage weights used to apportion the minimum wage effect would all surely be different if they could be defined for the teenage labor force. Incorrect specification of both the weights and hourly earnings series in the MWAGE variable casts doubt on the conclusion that the estimated coefficients are capturing the true effect of the minimum wage on the youth labor market.

Other Problems of Estimation

The minimum wage variable is not the only variable that suffers from poor data or problems in specification. The dependent variable, for example, has at least two such difficulties. First, as hinted at earlier, using the employment of "all teenagers" as a dependent variable biases the estimated disemployment effect of the minimum wage downward because the group is too inclusive. Many teenagers earn higher than minimum wages and thus are essentially unaffected by changes in the minimum. Furthermore, raising the minimum might result in employers hiring some previously part-time, unemployed, or nonparticipating high-productivity workers as a substitute for the now over-priced, low-productivity workers. In both cases, the inclusion of high-wage teenagers in the labor market being modeled would result in an underestimate of the decline in employment due to a change in the minimum wage! Second, labor market

data by race, especially unemployment statistics for minorities, are of very uneven quality. For this reason, the minimum wage effect estimated for minority groups tends to be much less reliable than that for whites or for the population as a whole.

Further difficulties also arise in the proper choice of the other independent variables, which are included in the equation to account for external conditions that might affect the estimated influence of the minimum wage. For example, if some variable serving as a proxy for labor market tightness were not included, movements in teenage employment due to changing, economic conditions might be mistakenly attributed to minimum wage changes.¹² On the other hand, data simply don't exist in usable form for variables such as family income and educational attainment, which almost certainly affect the supply of labor, and for variables like changing labor quality, which would be expected to affect demand. Instead, variables like these tend to be omitted from the equations that are estimated and their effects are incorrectly attributed to other variables in the model.

One specific controversy in the literature highlights the problems of appropriate variable choice. One group of studies finds evidence of significant unemployment effects of the minimum wage, while a second group finds little such evidence. While many differences exist in the two sets of studies, Michael C. Lovell has shown that the fundamental

¹² Note, however, that a variable such as the often-used unemployment rate of adult males may be a very imperfect control for labor market tightness as it pertains to teenage employment. The teenage industrial and occupational distributions are very different from that for adult workers. Because of this, shifts in demand for teenage workers may not correspond to overall demand shifts in the economy.

¹¹ Gramlich, p. 432.

cause of the conflicting conclusions is that those studies which find no significant unemployment effect include as one of their independent variables a variable to capture the rapid population growth of teenagers. Only when this growth variable is omitted is a significant unemployment effect found for the minimum wage. The argument for including the population variable is that it reflects the rapid increase in the supply of teenage workers which has tended to increase teenage unemployment. The arguments against including this variable are first, that the increasing supply of teenagers affects their unemployment only because the minimum wage keeps market wages from falling and thereby clearing the market, and second, that serious econometric problems arise if both population and minimum wage variables are included in the model. The question then is, should such a growth variable be included? As the case for both sides is strong, this important issue is still **unresolved**.¹³

SOME IMPORTANT NEW EVIDENCE

Largely because of shortcomings in the data, the final word on the impact of the minimum wage is clearly not yet in. Nevertheless, the quality of the studies examining this question has been improving as more refined estimation techniques are applied and new data

¹³ Articles by Fisher, and Douglas K. Adie and Lowell Galiaway conclude no: articles by Hyman Kaitz and Lovell say yes. For a summary of the controversy, see Goldfarb. See also Fisher, "The Minimum Wage and Teenage Unemployment: A Comment on the Literature;" Adie and Gallaway, "The Minimum Wage and Teenage Unemployment: A Comment;" and Lovell, "The Minimum Wage Reconsidered," all in *The Western Economic Journal*, December 1973. Also see Kaitz, "Experience of the Past: The National Minimum" in *Youth Unemployment and Minimum Wages*, Bulletin 1657, Bureau of Labor Statistics, 1970, pp. 30-54.

sources are utilized. Two such recent studies by **Ragan** and Gramlich provide added insight into this important **issue**.¹⁴

Ragan's Analysis

Using a two equation model, James F. **Ragan**, Jr., examines the effect of minimum wages on the youth labor market. His two dependent variables are the fraction of the teenage population employed and the fraction of the teenage population in the labor force. Using results from these two equations, he is able to construct estimates of unemployment rate elasticities, that is, the percentage response of unemployment rates to given percentage changes in the minimum wage. The minimum wage variable he uses is similar to the combined minimum wage/coverage variable discussed earlier (MWAGE), but has the additional refinement of weighting each industry by its share of youth employment rather than its share of total employment. In addition, through use of previously unexploited data, he is able to estimate separate equations for 16 population groups delineated by student and nonstudent status, age (16-17 and 18-19 years old), sex, and race.

Partly because of better data, **Ragan's** findings are more conclusive for men than for women and for whites than for blacks. His results, nevertheless, are broadly consistent, and point to a major disemployment effect for the minimum wage. For example, **Ragan** calculates the employment loss in 1972 from the 1966 minimum wage amendment which increased coverage and raised both the long-time covered minimum from \$1.25 to \$1.60, and the newly covered rate from \$1.00 to \$1.60. **Ragan's** model reveals that, had the amendment not been

¹⁴ **Ragan**, "Minimum Wages and the Youth Labor Market," *The Review of Economics and Statistics*, May 1977; and Gramlich.

enacted, total youth employment in 1972 would have been higher by 225,000, or 3.3 per cent; nonwhite employment 13.8 per cent higher, and white employment 2.4 per cent higher. Similarly, because of the amendment, the teenage unemployment rate rose 3.8 percentage points by 1972. This consisted of an increase for males of 4.5 percentage points; for females, 2.9 points; for whites, 3.9 points; and for nonwhites, 3.0 points above the level that would have existed had the amendment not been enacted.¹⁵

Ragan's research also contributes to the debate over whether it is appropriate to include an independent variable which captures the effect of teenage population growth on employment. On the hypothesis that the relative supply of teenagers affects their employment and participation, Ragan includes a teenage labor supply variable in both his equations in the form of the ratio of a teenage group's population to the total population 16 years of age and older. Unlike earlier studies, which used the unemployment rate as the dependent variable, Ragan's model with the relative supply variable shows the significant minimum wage disemployment effects reported above. Further disaggregating his sample, however, Ragan shows that it is only for nonwhite youth that their increased relative population has resulted in serious disemployment; white teenagers show no such effect.

Gramlich: The Impact of Minimum Wages on Employment and Income

The study by Edward M. Gramlich is perhaps the most comprehensive in the

¹⁵ Ragan explains the greater unemployment effect for whites than for nonwhites by another finding which shows that the higher minimum wage tends to slightly encourage white labor force participation, while it slightly discourages participation among nonwhites.

minimum wage literature—treating not only the disemployment question (how much employment falls in response to an increase in the minimum wage), but also issues such as compliance, coverage, the reaction of other wages to the changing minimum, and the inflationary impact of the minimum wage. Gramlich's central point is that whatever else the minimum wage is intended to do, it is basically an attempt to alter the distribution of income in favor of low-income families. The question is whether it effectively accomplishes that goal. Gramlich's approach is both innovative and extremely complex. Working from two theoretical models of the labor market, he derives criteria for assessing the conditions under which low-wage workers, as a group, would be made better off by increases in either the minimum wage level or its coverage.

Gramlich estimates separately the actual disemployment effect of the minimum wage for low-wage groups of teenagers, adult men, and adult women. He then calculates whether this disemployment is sufficiently large that each group of low-wage workers, taken as a whole, would perceive itself as worse off after the minimum wage change according to the criteria derived from his theoretical models. Using this method, Gramlich obtains somewhat ambiguous results. First of all, he finds disemployment effects for teenagers and adult men, but none for adult women. Nevertheless, he concludes that none of the groups are made worse off by a higher minimum wage because the disemployment effects do not outweigh the greater income received by those continuing to be employed at the higher minimum?'

¹⁶ Gramlich estimates that a 10 per cent increase in the minimum wage would yield a 0.9 per cent fall in the employment of teenagers. His estimates

Gramlich, however, goes on to show that the conclusion of no overall harm is quite deceiving for the teenage group. It turns out that, when each population group is considered in terms of the effect of the minimum wage on full-time and part-time employment, a striking result is obtained. As Gramlich observes:

"What is happening . . . is that high minimum wages reduce full-time employment of teenagers substantially, forcing many of them into part-time employment. The net result is the relatively slight overall disemployment effect typically found in other studies. If this is why disemployment is so slight, the most reasonable verdict is that teenagers have more to lose than to gain from higher minimum wages: they appear to be forced out of the better jobs, denied full-time work, and paid lower hourly wage rates; and all of these developments are probably detrimental to their income prospects in both the short and the long run."¹⁷

A similar analysis for adult men also shows a noticeable increase in part-time

are generally lower than those obtained in other studies of teenagers, which calculate approximately a 2 per cent fall in teenage employment following a 10 per cent increase in the minimum wage. Gramlich's conclusions on this point, however, depend crucially upon the plausibility of his model, upon the many assumptions he had to make regarding parameter values, and upon the generally poor quality of the data. Sharply different estimates of the disemployment elasticities and welfare loss might have been obtained using different, equally plausible, assumptions. See Robert J. Flanagan and Michael L. Wachter, "Comments and Discussion," BPEA, 2:1976, pp. 452-61.

employment due to the minimum wage, though this is hypothesized to affect mainly 20-25 year olds and those over 65. Results for adult females, on the other hand, show no such response. On the contrary, Gramlich concludes, a reasonable explanation of what is happening is that the higher minimum wage mainly benefits adult females who are attracted from the part-time into the full-time labor force, forcing low-wage teenagers into part-time employment.

The Minimum Wage, Inflation,
and the Distribution of Income

In his study, Gramlich also examined the impact of the minimum wage on inflation and the distribution of income. An increase in the minimum wage tends to raise the rate of inflation in two ways. First, because wages in the covered sector rise relative to the uncovered sector, minimum wage increases encourage unemployed workers to refuse uncovered employment while waiting for a job in the covered sector. As a result, the unemployment rate rises and the trade-off between unemployment and inflation tends to worsen? Second, the higher minimum wage tends to raise the total wage bill in the economy, both directly—because covered low-wage workers get higher wages—and indirectly—because other wages tend to rise when the minimum increases. For example, the 25 per cent increase in the

¹⁷ Gramlich, pp. 442-43. Whether the disemployment effects estimated in earlier studies are "slight" is, of course, a matter of opinion. Even Gramlich's small 0.09 estimated elasticity means that the 25 per cent minimum wage increase in 1974 lowered teenage employment by 2.3 per cent and raised the teenage unemployment rate by 2 percentage points (Gramlich, footnote 30). Note, also, that Gramlich's elasticity estimate for full-time teenage employment is 0.50, almost six times as large.

¹⁸ Wachter, p. 459.

minimum to \$2 in 1974 directly raised the wage bill by about 0.4 per cent, and indirectly raised it another 0.4 per cent.¹⁹ Such an increase in labor costs tends to be passed on by employers in the form of higher prices.

Gramlich's findings also reveal that the relationship between low-wage workers and low family income is very weak, and therefore, that the minimum wage has little positive effect upon the redistribution of income toward low-wage families. The results for teenagers were especially dramatic. Among families that contain teenagers who receive low wages, the median family income was \$12,900, substantially higher than the median income of families containing high-wage teenagers. Furthermore, fully 40 per cent of low-wage teenagers were members of families whose reported income exceeded \$15,000. From these and other data, Gramlich concludes, "The generally loose correlation between wages and family incomes implies that minimum wages will never have strong redistributive effects."²⁰

CONCLUSION

Analysis of the impact of minimum wage legislation on the labor market has been an extremely popular subject in the economic literature in recent years. The effect on youth unemployment has received particular attention for several reasons. First, young people are much more likely than adults to work in low-wage employment. Second, the youth unemployment rate, especially that for

minority teenagers, has risen dramatically in recent years. Third, this tremendous unemployment increase has occurred just as a new minimum wage bill has been passed by Congress.

Among economists, the issue is not whether the minimum wage has good or bad effects. Economic theory and virtually all studies of the issue are in agreement that, at best, the minimum wage is a highly inefficient tool for redistributing income. At worst, it is also a major cause of economic dislocation, distortion, and unemployment. The question that the literature has generally addressed, then, is just how serious are the various effects on the economy. For a variety of reasons, however, this question has proven extremely difficult to answer.

A major factor in this difficulty is certainly the poor quality of the available data. Many of the major series needed for a thorough analysis simply do not exist. Other series are too recent to be useful in time-series analysis, while still other data are of such dubious quality that their use leads to very tenuous results. A second factor is the complexity of the real world, always a problem in research of this nature. Finally, even without these difficulties, there would probably remain among the model builders substantial disagreement as to the exact form of the "correct" model, and the proper variables to include.

Despite these complications, the great weight of evidence is on the side of those who stress the negative aspects of the minimum wage. Two recent studies, one by James F. Ragan, Jr., and one by Edward M. Gramlich, provide further support to this negative view. Using previously unexploited data and a new model specification, Ragan finds that the minimum wage has had a major disemployment effect on teenagers, especially on black youth. In addition, the

¹⁹ Gramlich, pp. 426-30.

²⁰ Gramlich, pp. 443-49. Low-wage workers were those who, in 1973, usually received less than \$2 per hour. High-wage workers usually received more than \$4 per hour.

participation rate of minority teenagers is reduced by the minimum wage, further worsening their economic predicament.

In a very complicated article, Gramlich examines a number of important issues involving not only teenagers, but adult males and adult females as well. Gramlich, too, finds a disemployment effect for teenagers, though of smaller magnitude than **Ragan**. He also finds such an effect for adult males, but not for females. Using these results and some guidelines derived from a series of complex theoretical models, Gramlich concludes that the low-wage sectors of these three groups may not be worse off, taken as a whole, because the disemployment effects do not outweigh the greater income received by those continuing to be employed at the higher minimum. However, because of the many assumptions involved in **Gramlich's** methodology, other economists have questioned the **validity** of this conclusion.

Other results, however, are striking. Gramlich finds, for example, that despite what he characterizes as "slight displacement" for teenagers as a whole, large numbers of previously full-time teenage workers have been pushed into involuntary part-time employment by the minimum wage. The income and training losses due to this displacement are probably quite severe. Gramlich also calculates a significant inflationary effect of the minimum wage, principally through its impact on the total wage bill. Finally, Gramlich shows clearly that the goal of redistributing income through the minimum wage, or lifting low-income families out of poverty by raising the minimum, is illusory. Especially for teenagers, but for adults as well, the relationship between low individual wages and low family income is so slight that the minimum wage is very unlikely to have the hypothesized strong redistributive effects.

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Federal Reserve Bank of Kansas City
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Household Spending:

How Strong Will It Be? *By Dan M. Bechter and Jack L. Rutner*

Spending by households has played a key role in the current economic expansion. Not since the post-World War II boom in the private sector has such a large share of national output gone for personal consumption. The contribution of households to recent economic growth is even more significant when their investments in homes are counted along with their purchases of goods and services. Since the trough of the recession in early 1975, both categories of household spending have absorbed more than two-thirds of the country's production, as compared with less than two-thirds in the 11 preceding years. Because of its relatively large size, the household sector will continue to be a prime determinant of the rate at which the economy grows in the months ahead!

This article provides some insight into the probable strength of real household spending as the economy enters its fourth year of recovery. It is suggested here that durable goods and **housing**—historically the most volatile components of household purchases—will provide a

key to the strength of household spending in 1978. Although expenditures on new homes and durable goods make up only one-fifth of total household purchases, weaker growth of these components in 1978 would indicate that total household spending is also likely to grow more slowly. However, even if real spending by households on durables and homes does not grow at all, total household purchases are still expected to grow moderately in 1978 because the other components of household expenditures are likely to grow at about the 5 per cent rate projected for real disposable personal income.

THE HOUSEHOLD IN THE NATIONAL INCOME AND PRODUCT ACCOUNTS

The product side of the national income and product accounts (NIPA) focuses on the sources of final demand (the buyers of goods and services produced) during a calendar year or quarter (Table 1, left column). The logic is that one can measure the value of a period's production by tallying up expenditures on it. Domestic buyers of the nation's output are classified as either consumers, businesses, or government. Their respective NIPA expenditures are defined as personal consumption, gross private domestic investment, and government purchases. Foreign buyers of U.S. goods

¹ This article is concerned with the economic stimulus arising from personal expenditures—what people spend on themselves and their families. Data on such expenditures and financial flows make up the preponderant proportion of statistics on the household sector, which also include transactions involving personal trusts and nonprofit organizations serving individuals.

Table 1
THE HOUSEHOLD IN THE NATIONAL INCOME AND PRODUCT ACCOUNTS, 1976
(In Billions of Dollars)

<u>Product</u>		<u>Income</u>	
Gross National Product	\$1,707	Gross National Income	\$1,707
Personal Consumption Expenditures		Capital Consumption	179
Durable Goods	1,094	Indirect Business Taxes	163
Nondurable Goods	159	National Income	1,364
Services	443	Rental Income and Net Interest	112
Government Purchases	492	Corporate Profits	128
Gross Private Domestic Investment	361	Proprietors' Income	88
Fixed Investment	243	Compensation of Employees	1,036
Nonresidential	230		
Residential	162	Addenda:	
Household*	68	Personal (Household) Income	1,383
Nonhousehold	59	Rental Income and Net Interest	112
Changes in Business Inventories	9	Corporate Dividends	36
Exports	13	Proprietors' Income	88
Imports	163	Compensation of Employees <i>Less</i>	
	-155	Social Security Taxes	912
		Transfer Payments to	
		Persons, and Consumer	
		and Government Interest	235
Addenda:			
Household Expenditures	1,153	Personal (Household) Income	1,383
Personal Consumption	1,094	Disposable Personal Income	1,186
Residential Construction*	59	Personal Taxes	197

The household component of residential fixed investment is an estimate of the amount of investment in housing during the year by owners who occupy the new homes they buy or the existing homes they improve. The figure used here is the flow-of-funds estimate of residential investment by the household sector (\$57.6 billion in 1976), plus the flow-of-funds estimate of farm investment in residential construction (\$1.0 billion in 1976).

SOURCES: U.S. Department of Commerce and Board of Governors of the Federal Reserve System.

and services are all grouped under a fourth category—exports.

Two adjustments need to be made to the total of these categories because their sum does not quite equal total production in a particular period. The reason is that some goods purchased domestically during a particular period can be imports or could have been produced in earlier periods. Hence, in the first adjustment, the value of imports is subtracted from total purchases by subtracting it from

exports. In the second adjustment, the change in business inventories is added to investment, thereby taking account of the difference between current production that remains unsold and past production that is sold currently.

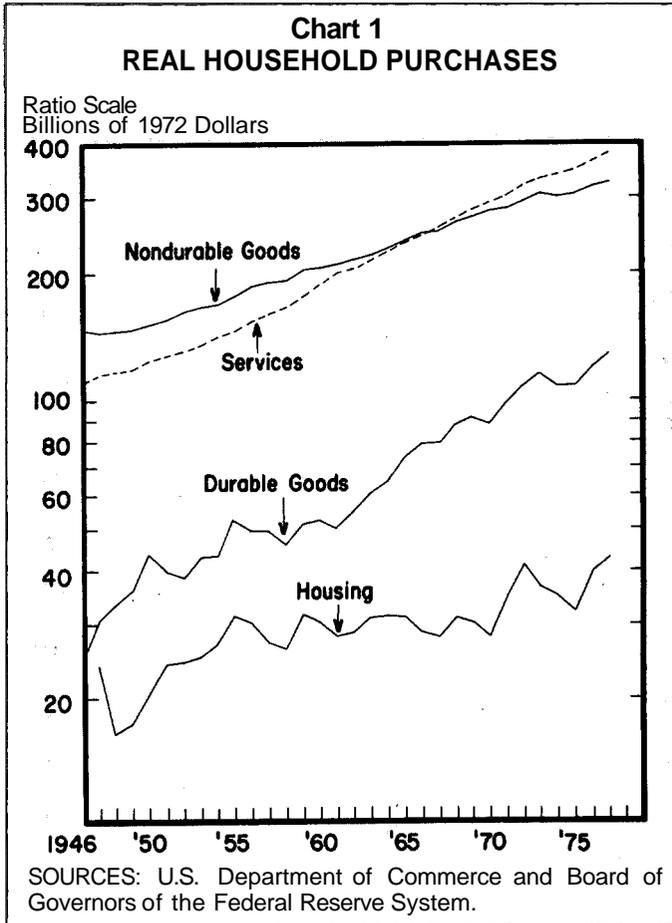
Household expenditures are included in both the personal consumption category and the gross private domestic investment category (Table 1, left column) of gross national product (GNP). As consumers, households make all

personal consumption expenditures and, as "businesses," households invest by buying new homes for their own use, as well as by improving the homes they own and occupy? Household income is derived from several sources identified on the income side of the NIP accounts (Table 1, right column), and is termed here personal income. Disposable personal income, or personal income minus personal taxes, is a measure of the household's ability to buy goods and services.

A GENERATION OF HOUSEHOLD PURCHASES

Changes in the growth rates of total household purchases during the past 30 years generally have resulted from recessions. As Chart 1 shows, real household expenditures on **durables** and housing generally reflect swings in economic activity. During recessions, real household purchases of nondurable goods and services usually do not decline because items such as food and rent are not postponable. However, purchases of new durable goods and homes are more easily postponed because households can continue to use existing stocks. During a recovery, individuals purchase the homes, cars, appliances, furniture, and other items deferred during the recession. This "acceleration" in **house-**

² Treating household expenditures that add to the nation's housing stock as investment, **while** treating all other household purchases **as** consumption, is one of the arbitrary characteristics of the NIP accounts. Ideally, perhaps, household expenditures on all durable goods, including housing, should be counted as investment. The using up of these goods (their depreciation) could then be considered part of **consumption**. This is the approach taken in the Federal Reserve Board's flow-of-funds accounts, which also go a step further than the NIP accounts by providing a measure of the household sector's **investment** in **residential construction**.



hold purchases diminishes, however, once stocks of durable goods and homes reach "desired levels."

As indicated, households purchase new durable goods and homes in order to bring their stocks of these assets more closely into line with levels they desire. Thus, the study of patterns in these purchases properly falls into the category of "stock adjustment" analysis. A stock adjustment model which fits the expenditure data fairly well assumes that purchases during any period are proportional to the difference between the actual stock and the desired stock. The following sections use this approach to determine how rapidly household

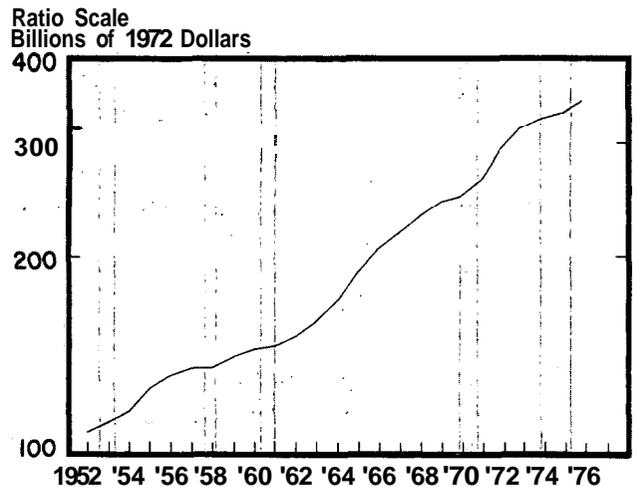
purchases will grow in the year ahead. Accordingly, attention is given in the following sections to developing quantitative estimates of actual and desired stocks of durable goods and housing.

Household Durables and the Stock-Adjustment Model

The value of the actual stock of consumer durables is a one-figure, dollar measure of the many kinds of durable goods owned by households. New goods can be valued at market prices. The values of used durables, however, are usually less than their original prices because of depreciation. If prices of used durables were readily available, as well as information on the number, types, and ages of used durables households actually own, the current value of the stock could be calculated. Such an approach can be followed for automobiles, using data on registrations and used-car prices by model, year, and make. Data on quantities and prices of other used consumer durables, however, are much less detailed.

The actual stock of household durables can be approximated by assuming the value of the stock increases by the amount of expenditures on new durables and decreases by some constant rate of depreciation on the stock of the previous period. New durable goods purchased during a particular period, such as a calendar quarter, are assumed to have been owned, on average, for one-half of that period, so a case can be made for depreciating them at one-half the full period rate. However, since a new durable good suffers sudden depreciation following its sale, a full period's rate of depreciation, the same as that applied to used durables, is applied to new durables. A slightly refined version of this method resulted in the data used in Chart 2. The annual rate of depreciation

Chart 2
STOCK OF CONSUMER DURABLES



SOURCE: Board of Governors of the Federal Reserve System.

NOTE: Shaded areas represent business cycle contractions as defined by the National Bureau of Economic Research; unshaded areas represent expansions.

assumed is about 29 per cent; constant dollar figures are used to remove the distortions caused by inflation?

The desired stock of durables cannot be measured directly. However, the stock of durable goods households want to hold can be assumed to depend upon household income. Specifically, it is assumed that households want to consume more goods, including durable goods, as their incomes increase?

Differences between actual and desired stocks can be assumed to average out to zero over long periods. This assumption allows using the long-term relationship of the actual stock of consumer durables to

³ The 29 per cent rate of depreciation in the real value of consumer durables is the rate used in the SSRC-MIT-PENN Quarterly Econometric Model of the U.S. Economy. It is believed this rate reflects market prices of used consumer durables.

⁴ This point was further developed by Dan M. Bechter in "Consumer Demand for Durable Goods," *Monthly Review*, Federal Reserve Bank of Kansas City, November 1974, p. 6, footnote 4.

Table 2
HOUSEHOLD PURCHASES OF DURABLE
GOODS, THE STOCK OF HOUSEHOLD
DURABLE GOODS, AND THEIR RATES
OF GROWTH
1970-76

	Billions of Constant 1972 Dollars		Per Cent	
	Stock at End of Year	Purchases During Year	Rate of Growth of Stock	Rate of Growth of Purchases
	(1)	(2)	(3)	(4)
1970	258.9	88.9	3.2*	-3.3*
1971	273.0	98.1	5.4	10.3
1972	294.5	111.3	7.9	13.5
1973	318.6	121.8	8.2	9.4
1974	328.9	112.5	3.2	-7.6
1975	334.0	112.7	1.6	0.2
1976	352.3	127.5	5.5	13.1
1977 est.	372.0	137.5	5.6	7.8
1978 forecast	396.5	145.0	6.6	5.5

*The 1969 values required to compute these percentage changes for 1970 are \$250.8 billion for the stock of household durables, and \$91.9 billion for purchases of household durables.

SOURCE: Federal Reserve Bank of Kansas City version of the SSRC-MIT-PENN Quarterly Econometric Model of the U.S. Economy.

income as an approximation of the relationship of the desired stock of durables to income. The long-term relationship can then be used to estimate short-term adjustments in the desired stock, given changes in real income. For example, if the actual stock of consumer durables has trended upward about 2 percentage points faster annually than household income, it may be inferred that, if income grows 5 per cent in the next year, the desired stock of durables will grow 7 per cent.

Before proceeding, it should be emphasized that the rate of growth of the stock of consumer durables differs from the rate of growth of additions to the stock of consumer durables (Table 2).

This article seeks insights into the probable strength of the latter—that is, into the rate of growth of household purchases of durable goods, which are additions to the stock of consumer durables. Such insights can be gained by noting discrepancies between rates of growth of the actual and the desired stocks. It should be kept in mind, though, that small changes in the rate of growth of the actual stock of durable goods may be associated with very large changes in the rate of growth of household purchases of durable goods. The past few years provide ample evidence of the lack of a simple correspondence between these rates of growth:

The data in Table 2 suggest that the large rate of growth of household purchases of durable goods in 1976 was primarily due to the recession-depressed level of purchases in 1974 and 1975. Whatever the reason for the 1976 rate of growth of household purchases of durable goods, sustaining that rate was required if the economy was to get the same boost from this source of demand in 1977. Judging from the first 3 quarters of data for 1977 and monthly data since, the rate of growth of household purchases of new durables is estimated to have been only 7.8 per cent in 1977, as compared with 13.1 per cent in 1976. The reason for the slower growth rate in 1977 can be understood, and the strength of household durables demand in 1978 can be forecast, by referring to the behavior of the desired stock of household durables.

As indicated earlier, the desired stock of household durables can be assumed to depend upon disposable income. This dependence is further assumed to be measured by the relationship of the actual stock of household durables to disposable income over long periods of time. Over the 1961-73 period, the actual stock in real terms grew about 1.6

percentage points faster than real disposable income?

From 1972 through 1976, real disposable income grew at an average annual rate of 2.7 per cent. Based on the relationship for the 1961-73 period, an annual rate of growth in the desired stock of 4.3 per cent ($2.7 + 1.6$) from 1972 to 1976 could justifiably be postulated. In fact, the actual real stock of durables rose at an average annual rate of 4.6 per cent over this period, suggesting that, for the 1972-76 period as a whole, actual and desired stocks grew about the same amount. This does not necessarily mean that actual and desired stocks were equal by the end of 1976. But it does imply that the discrepancies between the two that arose in individual years between 1972 and 1976 were reduced by the end of the period.

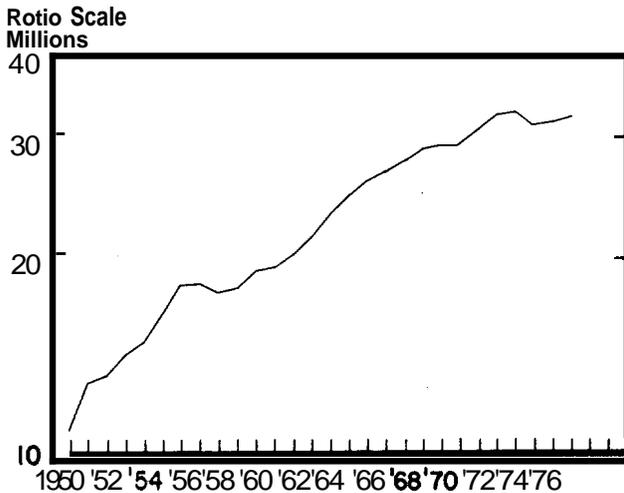
The slower growth rate of purchases of consumer durables in 1977 can now be better explained. During 1976, real disposable income rose 3.8 per cent, while the actual stock of durables rose 5.5 per cent, slightly more than the estimated increase in the desired stock of 5.4 per cent. But, as shown in Table 2, this 5.5 per cent increase in the stock was associated with a 13.1 per cent increase in purchases. Now, suppose real disposable income grew at a rate of about 4 per cent in 1977, which seems likely from preliminary data. Then the growth in the actual stock of durables necessary to maintain the relationship of actual and desired stocks would have to have been about 5.6 per cent—slightly more than the growth of the stock in 1976. A 5.6 per

⁵ The best candidate for explaining the "extra" 1.6 percentage points of growth is the declining relative price of durable goods. Between 1961 and 1973, the price index of consumer durable goods rose 26 per cent, as compared with a 37 per cent increase in the price index of nondurable goods, and a 63 per cent increase in that of services.

cent increase in the stock of consumer durables from the end of 1976 to the end of 1977 is equivalent to a net dollar increase of \$19.7 billion, from \$352.3 to \$372.0 billion (Table 2, column 1). This net increase of \$19.7 billion in the stock required a total of \$137.5 billion in purchases of new durables during 1977, of which \$117.8 billion offset depreciation. Thus, an increase of 7.8 per cent in purchases of new durables, from \$127.5 billion in 1976 to \$137.5 billion in 1977, kept the actual stock in line with the desired stock, given the estimated 4 per cent increase in real disposable income. The point to be emphasized is that a small increase in the rate of growth of the stock, from 5.5 per cent in 1976 to 5.6 per cent in 1977, was accompanied by a substantial decline in the rate of growth of new purchases, from 13.1 per cent in 1976 to 7.8 per cent in 1977.

In 1978, the rate of growth of real disposable income will again be the key to the strength of household spending. If the 1978 growth in real disposable income is close to the consensus forecast value of 5 per cent, the desired stock will increase, according to the analysis above, by $5.0 + 1.6 = 6.6$ per cent. If the actual stock of durables grows as much as the desired stock, its value at yearend 1978 will be \$396.5 billion (Table 2, column 1). This would be an increase of \$24.5 billion from the yearend 1977 value of \$372 billion. An increase of \$24.5 billion in the stock of consumer durables during 1978 would require purchases of new durables totaling \$145 billion, which is the value shown forecasted in Table 2. However, the increase in the amount of purchases of new durable goods from \$137.5 billion in 1977 to \$145 billion in 1978 would be only 5.5 per cent, down from 7.8 per cent in 1977. Thus, the 1978 increase in new durables purchased by households is not likely to be sustained at the 1977 rate,

Chart 3
STOCK OF AUTOMOBILES
IN NEW CAR EQUIVALENT UNITS
(Annually as of July 1)



SOURCE: Wards Automotive Yearbooks.

which itself was much slower than the rate of increase in 1976.

Before leaving the consumer durables category of household purchases, a subsection on new car purchases is included in recognition of the importance of the automobile in consumer demand and because of the superior data available.

New Automobiles: Much of the sensitivity of consumer durables expenditures to economic fluctuations is due to ups and downs in new car purchases. If further growth in new car sales is in prospect, the chances for sustaining the rate of increase of household spending are enhanced.

Household purchases of new cars are also a good example of the stock-adjustment process. Chart 3 graphs the stock of cars in new car equivalents over time. As was true of the stock of consumer durables—which was measured net of depreciation—the stock of new car

equivalents measures a net stock in new car equivalents which differ from ordinary units by an amount of depreciation. The gross stock, in contrast, is simply an unadjusted count of cars on the road. This gross stock (not charted) stood at 36 million in 1950 and had risen to nearly 100 million by 1977.

The net stock of autos could also be calculated in dollar value by using market prices for new and used cars, or by applying rates of depreciation to the original price of autos in operation. The new car equivalent method used to generate the stock for Chart 3 is much easier. A new car, regardless of value, is counted as one unit. The new car equivalent value of any other car in use is assumed to be 75 per cent of its new car equivalent value the year before—a double declining-balance method of depreciation. That is, after one year of use, a car counts as 0.75 units in new car equivalents, after two years it counts as $0.75 \times 0.75 = 0.5625$ units, etc.

The desired stock of autos can be estimated by the same approach used for estimating the desired stock of all consumer durables. Between 1966 and 1973, the actual stock of autos (Chart 3) grew on the average about three-fourths as fast as did real disposable personal income over that period⁶. If the annual growth in the desired stock of autos can be taken to be three-fourths of the annual rate of growth of real household income, the desired stock fell about 1½ per cent in 1974, rose about 2 per cent in both 1975 and 1976, and then rose by another 3 per cent in 1977.

If the desired stock did indeed follow the path just indicated, then its decline of

⁶ A somewhat shorter period is used here for autos than was used for all durables because the relationship between the average rate of growth of the auto stock and that of income seems to have changed since 1965.

Table 3
THE STOCK OF AUTOMOBILES AND NEW CAR SALES IN THE
UNITED STATES: UNITS AND PERCENTAGE CHANGES, 1973-78

Year	Stock of Passenger Cars in Thousands of New Car Equiva- lent Units as of July 1	Percentage Change in Stock of Cars	Retail Sales of New Passenger Cars in Thousands of Units, Year Ending June 30	Percentage Change in New Car Sales
	(1)	(2)	(3)	(4)
1973	32,800	-	11,739	-
1974	33,000	0.6	9,913	-15.6
1975	31,700	-3.9	8,322	-16.0
1976	32,000	0.9	9,714	16.7
1977	33,100	3.4	10,753	10.7
1978 forecast	34,300	3.6	11,160	3.8

1.5 per cent in 1974, as compared with an increase in the actual stock of 0.6 per cent, suggests that households were, at that time, overstocked with autos. A possible explanation for actual stocks rising in the face of a decline in desired stocks is that households did not expect their real disposable incomes to decline; and, only after income declined did the buildup of the auto stock seem too large. Finding themselves burdened with an excess stock of automobiles, households cut their new car purchases to a level below that of depreciation on the existing stock. Thus, by mid-1975, the actual stock of autos (in new car equivalents) had been reduced by 6 per cent from mid-1974. This reduction was more than enough to bring the actual stock into line with what might normally be estimated to be the desired stock. At that time, however, households most likely did not consider this large reduction to be an overcorrection, because of the added uncertainties associated then with energy availabilities.

As of mid-1977, the actual stock of autos is estimated to have been about 1 per cent above its 1973 value in new car equivalents, although real disposable income was 6.6 per cent above its

mid-1973 level. Unless households no longer wanted as large a stock relative to income as they desired in 1973, the actual stock must be considered to have been well below the desired stock in mid-1977. If, for simplification, the actual and desired stocks are taken to have been equal in 1973, then the actual stock in mid-1977 was about 4 per cent below the desired level, assuming the desired stock grows at three-fourths the rate of increase of real disposable income.

The above analysis does not prove that the actual stock is below the desired stock, but there is no evidence of overstocking of autos currently. The likelihood that the actual stock is not now greater than the desired stock, together with continued expected growth in real disposable income, suggests that the year ending June 30, 1978, will be another good year for new car sales, as far as level of sales is concerned. But the rate of growth in new car sales is likely to decline again in 1978 as it did in 1977 (Table 3). The forecast value for the stock of automobiles on July 1, 1978 (bottom row, column 1 of Table 3), was derived by assuming households will maintain the July 1, 1977, relationship of the actual stock to their desired stock. To do so, the

actual stock would have to grow at three-fourths the 5 per cent rate of growth forecast for real disposable income. This growth of 3.6 per cent in the stock of cars is calculated to require 11,160,000 new cars in 1978.⁷

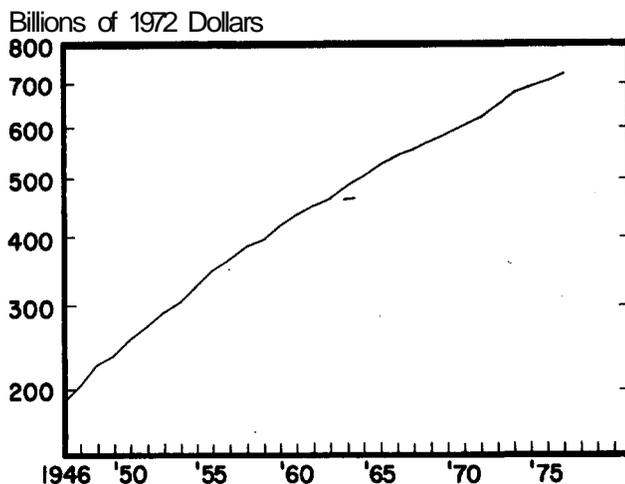
Owner-Occupied Housing

Homes last much longer than other household purchases. The slow rate of depreciation of the nation's stock of housing has been more than offset by gross investment in residential construction in every year since the end of World War II. To put it another way, the value of the net stock of housing in the United States has grown year after year for 30 years. But the rate of growth of this stock has been uneven because of the dramatic cycles in homebuilding. Stock adjustment analysis again proves useful in explaining the boom-to-bust behavior of residential construction, and therefore, in addressing the questions of the sustainability of that portion of residential construction activity attributable to household purchases.

Chart 4 shows a declining rate of increase in the real value of the net stock of owner-occupied housing during most of the past generation, although there was some acceleration during the 1971-73 housing boom. From 1966 to 1972, the net value of owner-occupied housing increased at an average annual rate of 3.2 per cent—virtually the same as the average annual rate of increase of real

⁷ New car sales of 11,160,000 between July 1, 1977, and June 30, 1978, would add an estimated 3,200,000 autos to the new car equivalent stock, according to the analysis in this article. The net addition to the stock is substantially less than new car sales because it will take 8,275,000 new cars just to offset the depreciation (25 per cent of 33,100,000) on the used cars from the 1977 stock that are still in use, plus 1,655,000 new cars to offset the losses, in new car equivalents, of automobiles retired from use for various reasons (estimated at 5 per cent of stock).

Chart 4
REAL NET STOCK OF
OWNER-OCCUPIED HOUSES



SOURCE: Department of Commerce.

disposable personal income over the same period. This relationship argues for using the annual rate of growth of real disposable income for the annual rate of increase in the desired stock of owner-occupied housing.

Real disposable personal income in 1976 was 11 per cent above its level in 1972. According to the assumed relationship between income and the desired stock of housing, the desired stock of housing at yearend 1976 was 11 per cent above its 1972 yearend level. The actual stock of housing at yearend 1976, in fact, was also 11 per cent above its 1972 level. These equal percentage increases in actual and desired stocks indicate that little, if any, gap remained between the absolute amounts of the two by the end of 1976, which would, other things equal, point to a decline in the rate of growth of household investment in housing in 1977.

Real household spending on housing is estimated to have grown 19.2 per cent in

Table 4
THE NET STOCK OF OWNER-OCCUPIED HOUSING, AND GROSS INVESTMENT BY HOUSEHOLDS IN RESIDENTIAL CONSTRUCTION, IN BILLIONS OF 1972 DOLLARS AND ANNUAL PERCENTAGE CHANGES 1972-78

Year	Stock	Per Cent Change	Investment	Per Cent Change
1972	\$655.7	—	\$41.3	—
1973	681.9	4.0	41.2	— 0.0
1974	697.8	2.3	35.7	—13.3
1975	709.3	1.6	33.0	— 7.6
1976	727.6	2.6	41.1	24.5
1977 est.	754.0	3.8	49.0	19.2
1978 forecast	784.0	4.0	53.0	10.0

Source of Actual Values: Department of Commerce.
Estimated and Forecast Values: See text.

1977, as compared with 24.5 per cent in 1976 (Table 4, column 4). The estimated \$49 billion households invested in homes in 1977 increased their stock of housing by 3.8 per cent, from \$727.6 billion to \$754 billion in 1972 dollars. Thus, only about half of the \$49 billion in residential construction for households took the form of an increase in stock, with the other half, or about 3 per cent of the previous year's stock, going to offset depreciation and other capital consumption of housing.

If real disposable income grows at a 5 per cent rate or more in 1978, it is conceivable that the rate of growth of household spending on housing could be maintained near its average for the past 2 years (about 22 per cent). A 5 per cent increase in real disposable income in 1978 would, according to the preceding analysis, increase the desired stock of housing by 5 per cent, to \$792 billion as of yearend 1978. This 5 per cent increase in the stock would require \$60 billion in gross investment in housing by

households in 1978 (\$22 billion in replacement + \$38 billion net increase). But there are several reasons for expecting the actual stock of homes to grow by less than 5 per cent in 1978 and, therefore, for forecasting a rate of increase of household spending on housing substantially less than 22 per cent, which is what an increase from \$49 to \$60 billion would mean.

Table 4 shows the stock of owner-occupied housing forecast to increase 4 per cent in 1978, which implies a 10 per cent increase in household investment in residential construction. (As indicated in the last row of Table 4, an increase of \$30 billion in net stock is estimated to require \$53 billion in gross investment by households.) While a 4 per cent increase in the actual stock is less than the 5 per cent increase forecast for the desired stock, this discrepancy is not a theoretical inconsistency. Households cannot always bring actual stocks in line with desired stocks quickly, as is the case when supply does not adjust immediately to demand. Such constraints would appear to apply to single-family homebuilding, and industry operating at peak rates in 1977. (A special factor constraining homebuilding in 1978 in some regions is a limit on natural gas hookups.) In 1973, for example, the third year of the previous housing boom, real disposable income rose 6.7 per cent while the housing stock rose but 4 per cent. Finally, during the current surge in homebuilding from 1975 to 1977, the net stock of owner-occupied housing has grown a total of 6.3 per cent, as compared with an 8 per cent increase in real disposable income. These considerations suggest a leveling of the rate of growth of the housing stock in 1978, which would mean a slowing in the rate of growth of household spending on residential construction.

Household Debt

Stocks of household durables and housing are assets that have their liability counterparts in types of household debt. Households often borrow on instalment plans to buy durable goods, and as the stock of household durables has grown, so has the amount of consumer instalment credit outstanding. Similarly, home mortgage debt owed by households has increased along with the stock of owner-occupied housing. Quite understandably, therefore, the prospects for sustaining household spending are closely related to the prospects for sustaining growth in household debt.

In real terms, both major classes of household debt have grown at diminishing rates since 1946. The outstanding amount of real consumer instalment debt grew at an average annual rate of 16.8 per cent between 1947 and 1956, 7.4 per cent between 1956 and 1965, and 4.3 per cent between 1965 and 1974. Real home mortgage debt has a similar history of growth in the postwar period. From 1947 to 1956, the amount of real mortgage debt households owed on their homes grew at an average annual rate of 12.7 per cent—dropping to 7.6 per cent in the 1956-66 decade, and then to a 3.4 per cent annual average from 1966 to 1976.

Interest payments are the primary burden of debt on households. As a ratio to disposable personal income, non-mortgage interest payments by consumers to business peaked at 2.4 per cent in 1965-66, and have stayed slightly below that percentage since. Although interest rates on some types of instalment loans have increased during the past 10 years, this has evidently been offset by a shift in the mix to less expensive types of instalment credit. The fact that non-mortgage interest payments have remained such a stable proportion of

household income suggests that, in the absence of a decline in interest rates, growth in consumer instalment debt will be held close to growth in household income.

Lack of data makes it difficult to assess the degree of strain from mortgage interest payments now felt by households relative to earlier periods. One estimate indicates that interest payments on home mortgage debt now require two and one-half times the share of disposable income required 20 years ago, and one and a half times the share of 10 years ago? Thus, in terms of interest payments relative to household income, the burden of home mortgage debt, unlike that of instalment debt, has increased sharply. This trend has negative implications not only for the growth of mortgage debt and

⁸ Two estimates of home mortgage interest payments were made for each year. Only one interest rate, that of FHA mortgages in the secondary market, was used for each year. For the "low" estimate of mortgage interest, it was assumed that the amount of home mortgage debt outstanding was always financed or refinanced at the lowest possible rates in the period. Thus, during periods of declining interest rates, refinancing at the new rate is assumed; during periods of rising rates, only additions to the amount of mortgage debt outstanding are assumed to carry the higher interest charges. For the "high" estimate of mortgage interest payments, just the opposite assumption is made: during periods of rising interest rates, all outstanding home mortgages are assumed to carry the most recent rate; during periods of declining rates, only additions to mortgage debt carry the recent market rate. According to the high estimate, mortgage interest payments by households rose from 1.5 per cent of disposable income in 1956 to 2.3 per cent in 1966 and to 4.3 per cent in 1976. According to the low estimate, the respective percentages are 1.3, 2.2, and 3.1. For the point made in this article, the increase in the burden of home mortgage interest is what is relevant, and this shows up clearly in either the high or the low estimate, or in the average of the two, which is what is referred to in the text.

new home purchases, but also for growth in other household spending.

SUMMARY AND CONCLUSION

Most of the nation's output—close to 68 per cent currently—is bought by individuals for personal or family use. Thus, demand by the household sector will be the key determinant of how fast the economy will grow in the year ahead.

An examination of types of household purchases shows that expenditures on durable goods and homes change most over time. The reason for the volatility of these components is that such purchases can be postponed as households continue to use carryover stocks. Measures of the actual stocks of consumer durables and housing prove useful, therefore, along with estimates of desired levels of these stocks, in analyzing household consumption and investment behavior. On the basis of this analysis, the current stocks of household durables and homes appear now to be close to desired levels. Accordingly, the rates of growth of household purchases of new durable goods and household investment in residential construction are likely to be somewhat slower in the current year as compared with earlier in the recovery.

Household instalment debt and home mortgage debt have risen along with household stocks of durable goods and housing. The rates of increase in these classes of household debt have slowed over the years, and now appear constrained to something near the rate of growth of disposable personal income. The interest payments on instalment debt have remained a relatively constant share of household income for many years, suggesting an implicit ceiling that will tend to prevent further large increases in consumer spending in excess of income gains. The interest payments on home mortgage debt, however, have grown as a share of household income, and while no ceiling on this proportion is yet in evidence, it is clear that this rising cost of shelter will curb income available for other purchases.

The weight of the evidence in this article points to a moderate rate of increase of household spending during the current year. The implication of this analysis is that if the economy as a whole is to achieve a real growth rate in the vicinity of 4½ to 5 per cent, which is the consensus forecast, sufficiently large increases in spending must occur in some of the other major sectors, such as in government purchases and business fixed investment.