

POPULATION AGING AND FINANCIAL MARKETS

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ABSTRACT

A number of financial market analysts have argued that the aging of the “Baby Boom” cohort contributed to the rise U.S. asset values during the 1990s, and that asset prices will decline when this group reaches retirement age and begins to draw down its wealth. Similar arguments can be raised for other nations that face analogous demographic transitions over the next half century. This paper explores the importance of changing demographic structure for asset returns, asset prices, and the composition of household balance sheets in the United States. The paper begins by discussing the theoretical basis for a link between asset markets and the age structure of the population. Standard models suggest that equilibrium returns on financial assets will vary in response to changes in population age structure, although the empirical magnitude of such effects is very difficult to predict with precision. While the direction of the effect of demographic changes is not controversial, the importance of such changes for financial markets is open to debate. The paper presents several strands of empirical evidence that bear on this issue. First, it reports an empirical analysis of the current age-specific patterns of asset holding in the United States, with the goal of understanding the age-wealth trajectory and how it may affect the future demand for financial assets. Data on age-wealth profiles from repeated cross-sections of the Survey of Consumer Finances suggest that asset holdings rise sharply when households are in their 30s and 40s. Aside from the automatic decline in the value of defined benefit pension assets as households age, however, other financial assets decline only gradually when households are in their retirement years. When these data are used to project asset demands in light of the future age structure of the U.S. population, they do not show a sharp decline in asset demand between 2020 and 2050. This finding calls into question the “asset market meltdown” view. Similar arguments are likely to apply in other nations where elderly households exhibit substantial saving rates. Second, the paper considers the historical association between population age structure and real returns on Treasury bills, long-term government bonds, and corporate stock. The evidence suggests only modest effects, if any, of demographics on returns. This is partly due to the limited power of statistical tests based on the few effective degrees of freedom in the historical record of age structure and asset returns. There is more evidence of a historical correlation between measures of asset levels, such as the price-dividend ratio, and summary measures of the population age structure, although once again the results are sensitive to choices about econometric specification. These empirical findings provide modest support, at best, for the view that asset prices could decline as the share of households over the age of 65 increases. Finally, the paper concludes with a discussion of the types of financial assets that will experience rising demand, and the types that will face declining demand, as the U.S. population and the population in other developed nations ages.

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The United States and many other developed nations are in the midst of a demographic transition. By 2030, the fraction of the U.S. population over the age of 65 will be greater than the current fraction over that age in Florida. This demographic transition will have profound social and economic implications. Its impact on government fiscal balance has received widespread and growing attention. The substantial rise in the ratio of retired workers to those in the labor force during the next three decades will place substantial strains on public pension and health insurance programs. Some combination of higher taxes, reduced program benefits, or protracted government deficits is a likely consequence. In many nations other developed nations, the fiscal prospect is even more daunting than it is in the United States.

Population aging is also likely to affect financial markets through channels other than government borrowing. Some argue that demographic changes, notably the entry of the Baby Boom cohort into their traditional high-saving years, contributed to the rise in stock prices during the 1990s. Others extend this argument to suggest that when the Baby Boom cohort reaches retirement, many households will be trying to sell financial assets to support retirement consumption, thereby driving down asset values. Such selling pressure could reduce the long-term rate of return earned by Baby Boom investors on their retirement saving. Siegel (1998, p.41) describes this concern: “The words “Sell? Sell to whom?” might haunt the baby boomers in the next century. Who are the buyers of the trillions of dollars of boomer assets? The [baby boomer generation] ... threatens to drown in financial assets. The consequences could be disastrous not only for the boomers’ retirement but also for the economic health of the entire population.” Schieber and Shoven (1997) develop this argument with regard to defined benefit pension funds: net flows into these funds will shift from positive to negative as the population ages. In addition to affecting the market-wide pattern of asset returns, aging populations may also change the composition of financial products demanded by the household sector.

The paper examines the potential impact of population aging on asset returns, the valuation of financial assets, and the demand for various financial products and services. It does not consider the implications of demographic shifts for government budgets or fiscal balance. It also sidesteps the important issues that surround the differential aging of different nations and the resulting international capital flows.

The paper is divided into seven sections. The first provides a brief review of the demographic changes that will confront the United States and other developed nations during the next half century. It summarizes the projected evolution of the share of the population between the ages of 40 through 64, and over the age of 65. Section two develops a conceptual framework for analyzing how an aging population, triggered by falling birth rates and rising life expectancy in old age, affects the demand for financial assets. A number of studies have modeled the impact of a “baby bust” on financial markets. This section describes the key features of the various models and it highlights modeling assumptions that have an important effect on the results. While these models suggest that demography should affect equilibrium asset returns, they do not offer precise guidance on the empirical magnitude of such effects. This makes the analysis of historical relationships, and cross-country patterns, essential. The third section reviews previous empirical research on the relationship between demographic structure and asset returns.

The next three sections present new empirical evidence on the links between population aging and financial markets. Section four investigates the age-specific pattern of asset holdings using data from the 2001 and earlier Surveys of Consumer Finances. It describes the challenges that arise in estimating how population aging will alter aggregate asset demand. It also develops a time series measure of projected asset demand based on age-specific patterns of asset demand in the 1990s, and future projections of population age structure. The fifth section presents new findings on the historical correlation between various measures of population age structure, asset returns on bonds and stocks, and the level of asset prices. As in the previous literature discussed in section three, the empirical findings are mixed. The most robust finding is a positive correlation between the share of the population in the prime working years and the level of stock prices, as measured by the price-dividend ratio. Even where statistically significant historical relationships emerge, however, there are often questions about the plausibility of the findings. In some cases the historical patterns, if extrapolated for the next three decades, suggest unreasonably large changes in asset prices. This raises the important possibility that demographic variables proxy for other omitted factors that may determine asset returns and asset prices. Minor changes in the econometric specification also appear to lead to substantial changes in the empirical results in many cases.

Section six moves beyond the analysis of aggregate financial market effects, and explores the potential effect of population aging on the demand for particular categories of financial assets. The results are not surprising: products such as annuities and long-term care insurance, which are demanded by households late in the life cycle, are likely to account for a growing share of financial market activity. There is likely to be greater demand for products that facilitate the preservation or the draw-down of wealth, rather than the accumulation of assets. A critical question in making quantitative projections about future asset demands is whether the age-specific pattern of asset ownership is stable over time. A brief conclusion outlines some of the unresolved issues that are likely to play an important role in determining the long-run effect of population aging on financial markets.

1. The Demographic Transition in the United States and Other Nations

Between 2000 and 2030, current projections suggest that the fraction of the U.S. population over the age of 65 will grow from 17 to 27 percent of the population over the age of 20. Figure 1 shows the historical and projected percentage over age 65 in the total U.S. population and in the adult (age 20+) population. Figure 2 presents analogous information on the population between the ages of 40 and 64. Table 1 reports the numerical data that underlying Figures 1 and 2, as well as additional information on the coming demographic shift. Table 1 shows, for example, that the median age of the U.S. population will rise by nearly four years over the next three decades.

Many discussions of the prospective impact of demographic change on aggregate financial markets emphasize the changing share of the population in the “asset accumulating years,” ages 40-64. Figure 3 provides some insight into the basis for this attention. During the last fifty years, the real level of stock prices, as measured by the real value of the Standard and Poor’s 500, has moved in tandem with the fraction of the adult population in this age group. Whether this pattern reflects a causal link between demographic variation and the stock market, and whether it is likely to predict future asset price movements, is the central issue that the empirical work presented below will address.

Table 1 describes the prospective changes in the share of the adult population between the ages of 40 and 64, and places these changes in historical perspective. The percentage of the total population between 40 and 64 rose by roughly four percentage points, to 30.1, between 1970 and 2000. It is projected to rise to 33.1 percent in 2010, and then to drop nearly five percentage points to 28.3 percent by 2040. Given the historical movements in this population share, it is important to recognize that the absolute size of the predicted change during the next half century is not substantially outside the range of historical experience. What is different, however, is the growing share of older individuals relative to children in the dependent population.

The penultimate column of Table 1 tracks the share of the over-20 population in the key asset accumulating years of 40-64. Individuals in this age group accounted for 42.2 percent of the adult population in 2000, an increase of six percentage points from just a decade earlier. This age group is projected to account for more than 45 percent of the adult population in 2010, but to decline to 38.3 percent of the adult population by 2050. Once again, however, the historical record shows other periods of substantial variation in this ratio. In 1960, 43 percent of the adult population was between the ages of 40 and 64. The median age of the over-20 population in 1960 was 45.3 years. In 2000, it was 46.7, and it is projected to rise to 52.5 years by 2050.

Populations can grow older for three non-exclusive reasons: a decline in the birth rate, a decline in the mortality rate at old ages, or a decline in the number of young immigrants. The projected changes in the age structure of the U.S. population are due to a combination of falling fertility and rising old-age longevity. The birth rate in the United States declined from 3.03 in 1950, to 2.43 in 1970, to slightly greater than 2.0 today. At the same time life expectancy for men at age 65 has increased from 12.8 years in 1950, to 15 years in 1990, to 16 years today. For women, the increase over the last half century has been even larger, from 15.1 years in 1950, to 19 years today. When a population ages because its existing old live longer, it will be challenged to transfer resources to individuals who did not expect to outlive their savings, but did so because of mortality improvements. When a decline in birth rates is the predominant source of population aging, there is more time to prepare for the older population. The current policy debates about

government transfer programs, and the discussions about how financial markets will be affected by prospective demographic shifts, illustrate precisely this pattern.

Demographic projections are more reliable than many types of long-term economic forecasts, because predicting the future number of individuals of a given age depends largely on the current number of younger individuals and the mortality rate over the prediction interval. It is nevertheless important to recognize two sources of potential uncertainty in these forecasts. The first is potential changes in either birth rates or mortality rates during the forecast period. The forecasts developed by the U.S. Census Bureau and the Social Security Administration do not project widely varying birth rates in the future, although there has been wide variation in the U.S. birth rate during the last half century. They also embody future improvements in mortality at a relatively constant rate. It is possible that new diseases or major medical innovations could have a dramatic effect on the future course of mortality rates. The greatest source of uncertainty in demographic forecasts, however, is probably immigration. If the U.S. were to substantially expand the number of immigrants who were allowed to enter the country over the next three decades, the rate of population aging would be slower than the data in Table 1 suggest. This is because the average age of immigrants is lower than the average age in the existing population. One open question about immigrants is whether they will decide to return to their home country when they reach retirement age. Such emigration could magnify the impact of immigration in offsetting the aging of the U.S. population.

The demographic shift facing the United States over the next few decades is matched, and in many cases exceeded, by the demographic changes in other nations. The decline in the U.S. birth rate during the last three decades has been smaller than that in many other developed nations, so the degree of population aging is somewhat attenuated. The pattern of aging in other developed nations is important for analyzing how financial markets in one nation, such as the United States, will respond to an aging society. When only one nation experiences a demographic transition, international capital flows can blunt the impact on financial markets and on real activity. When many nations experience a demographic transition together, however, this mechanism is no longer available to moderate the impact of the aging population.

Table 2 shows the evolution of the aged dependency ratio, the ratio of the number of individuals over the age of 65 to the number between the ages of 20 and 64, for a sample of developed nations. It also reports the total dependency ratio, which is the number of individuals over 65 or under 20 divided by the population aged 20-64. The table shows that the demographic transitions in several other developed nations are even more dramatic than the demographic transition in the United States. In Italy, for example, where the birth rate has fallen to well below replacement rate in the last decade, the aged dependency ratio is projected to rise from 29 percent to 51 percent between 2000 and 2030. The corresponding change in Japan is nearly as large: 28 percent to 56 percent. By comparison, the changes in the United Kingdom and Canada are closer to those in the United States, which rises from 21 percent to 37 percent.

2. Conceptual Analysis of Age Structure, Asset Prices, and Asset Returns

A variety of economic models can be developed to lend evaluate popular claims that demographic factors may have contributed to recent movements in asset values, and that they might have an important influence in the future. The challenge is to move beyond the simple intuition that “demography matters” and to develop insight on the potential magnitude of demographic effects, while also understanding the factors that are likely to magnify or attenuate the impact of demographic influences on asset prices.

A very simple model, sketched in Poterba (2001), can provide a starting point for understanding why demographic shocks may affect asset prices and asset returns. The model assumes that individuals live for two periods. They work when young (y) and retire when old (o). Normalize their production while working to one unit of a numeraire good, and assume that there is also a durable capital good that does not depreciate and that is in fixed supply. If the saving rate out of labor income is fixed at s for young workers, then demand for assets in a period when there are N_y workers will be N_y*s . With a fixed supply of durable assets (K), the relative price of these assets in terms of the numeraire good (p) must satisfy $p*K = N_y*s$. A “baby boom” which increases the size of the young worker cohort drives up asset prices so that the fixed physical supply of capital can meet the greater demand for financial asset holding. If a large birth cohort is followed by a small one, asset prices will increase and then decline. The return

on investments by the large birth cohort will be low, since this cohort will purchase assets at high prices. A small cohort, in contrast, will earn a high return.

This simple model neglects many important realities of asset pricing. Four of the most important omissions are the following:

(i) Fixed saving rate for young workers. A more sophisticated analysis would allow workers to vary their saving rates in response to expectations about future rates of return. This requires an optimizing model of household behavior that endogenizes consumption decisions. If changing age structure affects the prices of financial assets, the associated changes in rates of return may affect saving decisions. If the saving rate of workers in a large cohort was lower than that for workers in a small cohort, the resulting demand for capital would be smaller and the price of capital would be bid up by less than the simple model suggests.

(ii) Fixed supply of capital. Fixing the supply of capital amplifies the impact of shocks to asset demand. In a more realistic setting the price of capital goods would affect the growth of the capital stock. Abel (2001) and Lim and Weil (2004) show that allowing for a supply curve for capital goods can have an important impact on the link from demography to asset prices. In the extreme case of no adjustment costs to varying the capital stock, capital would always be priced at its reproduction cost. Demographic shocks could not affect the price of assets. This polar case is unlikely to be satisfied in practice, however.

(iii) Closed Economy without International Capital Flows. When the supply of capital in a single economy must equal the contemporaneous demand for that capital, intertemporal movements in the price of capital goods are likely to be larger than when international capital flows allow for a more elastic supply of capital. If all global capital markets were fully integrated, asset prices and rates of return would depend only on global demographic forces to the extent that they affected the supply of saving. The large gross flows of financial capital across borders make the closed-economy assumption untenable, but perfect capital market integration also seems inconsistent with the empirical evidence. While the strong correlation between a country's investment rate and its saving rate, documented by Feldstein and Horioka

(1980), appears to have weakened over time, it has not vanished. Obstfeld and Rogoff (2001) provide a recent overview of the evidence on capital market integration, and related puzzles such as the home bias in corporate equity investments. Prospectively, the integration of capital markets in currently emerging economies with those of developed nations may be an important factor determining the demand for financial assets.

(iv) Other Economic Effects of Population Aging. The foregoing analysis does not consider how a changing age structure might affect non-financial aspects of the economy, such as the rate of productivity growth. Factors such as the expected rate of growth play a central role in determining asset values and rates of return. Cutler, et al. (1990) suggest that links between age structure and the rate of productivity improvement, if they exist, can swamp many other channels linking demographic change to equilibrium factor returns. Bosworth, Bryant, and Burtless (2004) summarize existing evidence on how worker productivity varies over the lifecycle, and discuss a number of other channels through which a demographic shock may influence economic activity.

A number of research studies have explored the effect of population aging in stylized models that try to incorporate a more realistic description of saving behavior and asset price determination. The leading models focus on the closed economy setting, and they do not allow an aging workforce to affect the capital market through channels other than the supply of savings. They relax the stylized assumption of a fixed saving rate, and replace it with an overlapping-generations environment in which consumers live for many periods and formulate rational life-cycle plans. Some models also allow for a variable supply of capital, with adjustment costs.

Abel (2001, 2003) presents analytical results based on an overlapping generation model with a variable supply of capital. He shows that a stylized “baby boom,” in which the birth rate rises and then falls, would reduce the rate of return relative to what it would be in a steady state economy with a stable birth rate. Those born into a “baby boom cohort” would therefore face less attractive capital market opportunities than those born at other times. Abel (2001) also explores the sensitivity of his findings to alternative models of saving behavior, and in particular the impact of allowing individuals to have a

bequest motive. The basic results that emerge in models without a bequest motive can also obtain in models with a bequest motive, but the findings are sensitive to the specification of the bequest motive.

Several other studies have used calibrated versions of a numerical intertemporal general equilibrium model, again with hypothetical life-cycle consumers who “live” for many periods, to study how changing cohort size affects asset prices and asset returns. Three examples of such models, in chronological order, are Yoo (1994a), Brooks (2002), and Geanakoplos, Magill, and Quinzii (2004). All of these models suggest that a demographic transition will affect capital market returns, although the magnitude of the estimated effect varies across models.

Yoo (1994a) calibrates a model in which overlapping generations of consumers live for 55 periods and work for 45. He finds that a rise in the birth rate, followed by a decline, first raises then lowers asset prices. While this broad pattern is consistent with the claim that the baby boom cohort may face lower financial market returns over their lifetime, the effects appear to be quite sensitive to whether or not capital is in variable supply. With a fixed supply of durable assets, asset prices in the “baby boom economy” rise nearly 35 percent above their level in the baseline case. This effect is attenuated, to a 15 percent increase in asset prices, when capital is in variable supply. In the case of variable asset supply, the return on capital varies by 40 basis points in a simulation of a “baby boom” that is loosely calibrated to resemble that in the United States during the last four decades.

Brooks (2002) also presents simulation evidence in an overlapping generations economy. Unlike Yoo’s (1994a) specification in which individuals live for 55 years, Brooks (2002) assumes that individuals live for four periods. His model incorporates both risky and riskless assets, however, so it is possible to explore how demographic shocks affect the risk premium. The model is calibrated so that older individuals prefer to hold less risky assets. Rapid population growth that persists for half a generation (two periods) and which is followed by below-average population growth affects the equilibrium level of both risky and riskless asset returns. Equilibrium returns on the risky asset change by roughly half as much as the riskless return, so the equilibrium equity risk premium declines in the early stage of the “baby boom,” and then increases when the large cohort is old. Brooks’ (2002) simulation of a

stylized “baby boom,” again designed to mimic recent U.S. history, suggests that riskless returns change by about 30 basis points as a result of the demographic shift, while asset prices vary by less than seven percent as a result of this demographic shock.

Geanakoplos, Magill, and Quinzii (GKQ) (2004) develop an even more elaborate overlapping generations model in which they incorporate a number of factors, such as realistic age-income patterns, that improve the model’s similarity to the post-war U.S. economy. They also explore the sensitivity of their findings to allowing for Social Security, bequests, and a range of other factors. Their core findings suggest that demographic shocks like those experienced in the post-war United States could generate substantial swings in asset values, but that actual peak-to-trough movements in the stock market are two to three times greater than the demographic analysis can explain. On balance, the GMQ results suggest larger effects of demography on financial market returns, and on asset values, than either of the previous studies. They also develop subtle predictions about the comovement of riskless returns, the risk premium, and the value of claims on risky assets. The authors caution that the analysis is based on a closed economy model, and that international capital flows might moderate the effects.

One common feature of the simulation models described above is their assumption that agents have perfect foresight about demographic shocks. This implies that when a demographic shift such as a decline in the birth rate occurs, it begins to affect the asset market immediately. The long resulting lead times associated with demographic shocks mitigate the impact of these shocks on asset prices, because any potential adjustments such as changes in the capital stock can take place in advance. Whether capital market participants are in fact so far-sighted is an important factor in determining the impact of demographic change. An intriguing recent study by Della Vigna and Pollett (2003) raises some questions about whether capital market participants are fully forward looking. The study suggests that changes in demographic factors more than six years into the future do not have large effects on asset prices. Given the substantial popular literature on long-term demographic trends and asset markets, the strict assumption that demographic shifts are not anticipated seems untenable. However, if some market participants do not consider these trends in their

investment decisions, then the impact of future demographic change could be even larger than the foregoing models suggest.

The simulation models described above suggest that large demographic shocks have the potential to influence asset prices and asset returns, although the precise magnitude of these effects is not clear. In most settings, a plausibly-calibrated shock that resembles the baby boom in the United States appears to have a modest impact on asset returns. The results are sensitive to a number of modeling assumptions and choices about parameters, however, which motivates the analysis of historical data on asset returns, asset prices, and demographic structure. The next section provides a brief review of the empirical literature that has explored these issues. Unfortunately, the empirical work does not resolve many of the outstanding questions, so the guidance offered by theoretical models must play an important part in evaluating the likely effects of demographic change.

3. Existing Empirical Work on Demographic Structure and Asset Returns

Nearly a dozen studies have investigated the correlation between the prices of financial assets, or the returns on these assets, and demographic structure. Most of the research has analyzed time series data for the United States, but several studies have also explored the patterns in other nations. The results are mixed, with some studies finding what appear to be large effects of demographic structure, while other studies fail to reject the null hypothesis that population age structure and returns or asset prices are unrelated. The differences in the findings in the various studies can be attributed to differences in the econometric specification and in sample period. This section offers a brief review of the empirical literature, starting with the research on the United States and then describing the smaller set of studies on other nations. It highlights the limited amount of data that can be brought to bear to study demographic linkages to asset markets.

To provide some perspective on the increasing sophistication of the empirical literature that considers how demography affects asset returns, I will describe a number of notable studies in

chronological order. I begin with studies that focus on the relationship between asset returns and demographic variables, and then describe several studies that focus on other variables.

The first study of population age structure and financial market returns, by Bakshi and Chen (1994), began from the presumption that older individuals are more risk averse than younger ones. A rise in the average age of the population would therefore be associated with an increase in aggregate risk aversion, while all else equal would translate into an increase in the required risk premium in financial markets. This study found that the fit of an empirical Euler equation for the intertemporal variation in aggregate consumption could be improved if the parameter describing aggregate risk aversion varied with the average age of the U.S. population. The results imply that a demographic transition like that expected over the next three decades could have a substantial effect on asset returns. This analysis, however, is conditional on a host of maintained assumptions about consumer preferences, the absence of constraints on consumption behavior, and the age profile of risk aversion. Poterba (2001) notes that survey evidence on household risk tolerance does not provide strong support for the assumption that risk aversion rises with age.

Yoo (1994b) and Macunovich (1997) allow for a more flexible relationship between population age structure and asset demands. Yoo (1994b) finds that a higher fraction of the population in the prime saving years is associated with a lower real return on Treasury bills. Large standard errors make it impossible to draw firm inferences about the link between demographic structure and returns on longer-maturity assets. Even for Treasury bills, however, the results are quantitatively small. Macunovich (1997) uses a richer set of demographic variables to explain the postwar fluctuations in the real return on the Dow Jones Industrial Average. Her specification includes so many demographic measures, however, that she may be overfitting the variation in the data sample. The out-of-sample predictions from her models are unstable, and they often imply empirical effects of population aging that are several times larger than those in the simulation models discussed in the last section.

Bergantino (1998) follows a more sophisticated strategy, based on Mankiw and Weil's (1989) analysis of demographic factors in housing markets, to analyze how demography and asset markets

interact. He uses cross-sectional data from the Survey of Consumer Finances to estimate age-specific demands for corporate stock and owner-occupied real estate, and combines these estimates with data on the changing age composition of the population to create measures of aggregate demand for both corporate equity and housing. He finds a positive association between his measure of asset demand and the level of stock prices. These effects are clearest when he focuses on the low-frequency variation in the demographic demand variable. Bergantino (1998) concludes that demographic changes can explain a substantial share of the post-war fluctuations in equity prices. Moreover, the demographic changes that are projected for the next three decades would have a sizable effect on asset values if past patterns continue to hold.

Poterba (2001) builds on the earlier studies and re-examines the relationship between real returns on Treasury bills, government bonds, and corporate stock, and several indicators of demographic structure. The study emphasizes the limited number of effective degrees of freedom in time series studies of returns and demographic change. Since there is only one Baby Boom in the United States, it may be misleading to suggest that there are many years of data on demographics and asset market returns. It may be more accurate to view the existing data as the result of one realization on time-varying birth rates. Poterba's (2001) econometric results provide very limited support for a link between asset market returns and demographic variables. There is weak evidence linking population age structure to real returns on Treasury bills, as in Yoo (1994b), but no other clear patterns. There is some support for a link between price-dividend ratios and demographic variables. This result is similar to, although not as strong as, the finding in Bergantino (1998). Updated estimates of the core econometric models in Poterba (2001) are presented later in this paper.

Geanakoplos, Magill, and Quinzii (2004) present empirical evidence that complements their simulation analysis. Their results suggest that the real level of share prices, measured by the S&P 500 index, is related to the ratio of middle-aged to young individuals in the population. This "MY ratio" is defined as the number of 40-49 year olds divided by the number of 20-29 year olds. The results suggest a

statistically significant link between “MY” and real stock returns, with a change like that projected for the 2000-2050 period resulting in roughly a sixty basis point decline in real returns.

While most studies of demographics and financial markets have analyzed data on asset returns, Goyal (2003) investigates whether age structure affects the demand for cash payouts from the corporate sector. He analyzes the correlation between age structure and the sum of dividends and net share repurchases, and finds that an increase in the fraction of the population in the retirement years is associated with an increase in net payouts from the corporate sector, as well as a decline in the equity premium. Although the results are statistically significant, the study also considers prospective changes in population age structure, and concludes that such changes are likely to have at most a modest impact on asset returns.

In addition to the foregoing studies that focus on the United States, a number of studies have now used data from nations to explore related issues. Erb, Harvey, and Viskanta (1997) focus on the 1970-1995 period for a sample of developed and developing countries. They find a positive relationship in both developed and developing countries between stock returns and the change in the average age of a country’s inhabitants. Taken at face value, the results would seem to suggest that an aging population is associated with rising stock values. The interpretation of this finding is especially difficult, however, because of the many sources of variation in population age structure across nations. In many developing nations, for example, average age may proxy for changes in underlying economic conditions that reduce morbidity and mortality. It is not clear whether such demographic changes should be viewed as the driving force behind asset market movements, or whether they in turn reflect other factors at work in developing nations.

Brooks (1998) also exploits cross-country variation by relating the level of real equity prices to a demographic structure variable. This variable equals the ratio of the population aged 40-64 to the population older or younger than this group. For eleven of fourteen countries in the sample, there is a positive relationship between this demographic variable and the real stock price. A key question in evaluating these results is how to normalize share prices to account for differences in debt and other

forms of leverage that differ across nations. Using domestic demographic data to study domestic returns in small nations may also face empirical difficulties. In countries with a substantial share of foreign investors in their equity markets, such as Denmark, Belgium, and the Netherlands, it is unclear whether domestic demographic variables should have much impact on asset returns and asset values.

Davis and Li (2003) focus on a smaller sample of seven countries with substantial equity markets. They find a statistically significant effect of the share of the population between the ages of 40 and 64 on the level of real stock prices and on bond prices as well. Their study moves beyond most of the previous work in including control variables for non-demographic factors that may affect asset prices, such as the rate of economic growth, the inflation rate, and the recent volatility of the equity market. The findings are robust to the inclusion of these control variables.

Geanakoplos, Magill, and Quinzii (2004) also report some international evidence on the association between the “MY” ratio and real stock returns. They study France, Germany, Japan, and the United Kingdom, and find mixed results. For France and Japan there appears to be a link between the MY ratio and the real price of corporate equities, but the relationship does not emerge in the other nations.

The studies described above broadly suggest that demographic factors are correlated with the level of asset prices, although each empirical specification is open to some question. The findings appear to be sensitive to minor changes in the definition of the dependent variables or the structure of the econometric model. The search for robust relationships between returns and demographic variables is therefore ongoing. The next two sections present new empirical evidence on this issue.

4. Age Patterns in Asset Ownership in the United States

In the multi-period lifecycle models that underlie the overlapping generations models described above, household financial asset holdings evolve predictably over the life course, with accumulation while working followed by decumulation while retired. Yet Hurd's (1990) review suggests that empirical age-wealth profiles often display little if any decumulation in post-retirement years. This section begins by presenting data on age-specific asset holdings in the 2001 Survey of Consumer Finances (SCF), and it

then develops a measure of the change in asset demand between two dates that can be attributed to demographic change.

The SCF provides the most comprehensive information on asset ownership in the United States. The Federal Reserve Board commissioned the first "modern" SCF in 1983, and the survey has been carried out every three years since then. Aizcorbe, Kennickell, and Moore (2003) provide a detailed description of the 2001 SCF, along with information on the composition and concentration of household wealth in 2001. SCF data can be used to calculate average asset holdings for families headed by individuals of different ages. Table 3 presents information on mean and median net financial assets and net worth for the 2001 SCF. Net worth is a broader asset concept than net financial assets. It equals net financial assets plus holdings of both owner-occupied and investment real estate, less mortgage debt, along with vehicles, business equity, and miscellaneous other assets.

Table 3 presents both age-specific means and medians for net worth and net financial assets. At every age, the mean is several times greater than the median, reflecting the very substantial dispersion in private wealth holdings. Net financial assets are more dispersed than net worth. For households in their early 60s, mean net financial wealth is eight times greater than median net financial wealth. The substantial dispersion of asset holdings leads to standard errors of the means that are large enough to make it difficult to reject most hypotheses about the slope of the age-wealth profile at advanced ages. The substantial dispersion also implies that studies of how demographic change may affect asset demand should pay close attention to the lifecycle patterns of high net worth households, which account for a very large share of aggregate financial asset holdings.

The summary statistics in Table 3 suggest that average holdings of net financial assets and average net worth rise with the age of the family head until roughly age 60. Neither net worth nor net financial assets increase after age 60, but given the imprecision of the estimated age-specific means, it is not possible to reject the null hypothesis that the averages are constant beyond age 60. For net financial assets there is virtually no decline in old age, while for net worth the mean for families headed by someone over the age of 75 is below that for younger families, although the standard error is large. The

point estimate of the net worth level for individuals aged 75+ is roughly one quarter lower than that for households in their mid-60s.

Large standard errors notwithstanding, the point estimates of the age-specific means in Table 3, which suggest a limited decline in financial asset holdings as families age, suggests caution in concluding that there will be a rush to sell financial assets when Baby Boomers reach their late 60s and early 70s. Sabelhaus and Pence (1999) have carried out related analysis using the SCF data, and they find more evidence of asset draw-down in old age. The difference between their results and those in Table 3 is partly due to their use of earlier SCF data, but more importantly due to corrections that they make for age-related mortality differences. They observe that wealth and mortality are inversely related, so that the sample of families that survive to extreme old age is likely to be a wealthier group than the general population at an earlier age. Their procedure suggests that there may be somewhat greater decumulation than Table 3 suggests.

Table 3 presents cross-sectional age-wealth profiles from a single survey. Prior research on wealth accumulation emphasizes, however, that cross-sectional patterns may not describe the trajectory that a given cohort will follow as it ages. Shorrocks (1975) is one of the first studies to recognize the need to move beyond cross-sectional data in studying age wealth profiles. The asset holdings by an individual of age a in time period t can be decomposed into an age effect, a time-period specific effect, and a cohort effect for those who were born in period $t - a$. The age effect captures the effect of the “point in the lifecycle” on wealth holdings. The time effect recognizes the impact of the particular moment at which the survey is taken. After several years of favorable stock market returns, for example, average wealth at all ages will be higher than if the stock market has just fallen sharply. The cohort effect reflects life-long effects of date-of-birth. For example, individuals born prior to the Great Depression may have a greater desire to save than those born later, reflecting their greater experience with economic hardship and the loss of financial wealth. They also may have a smaller endowment of human capital as a result of their difficult early experiences in the labor market.

With a single cross-sectional data set, however, it is not possible to separate age, cohort, and time effects. With panel data or repeated cross-sections, it is possible to estimate two of the three effects, but it is not possible to recover all three. This is because the cohort effect is a linear combination of the age and time effects. This fundamental identification problem raises an important challenge for studying how population aging will affect asset demand. Two examples illustrate the difficulties. First, if older cohorts have lower lifetime earnings than younger cohorts, and if the accumulation of financial assets is correlated with lifetime earnings, then one might observe lower wealth levels for those at advanced ages but this need not be the result of decumulation. Second, if the asset market returns over the life of one cohort were more favorable than those over the lifetime of another cohort, the wealth at any age of the first cohort might be higher than that of the second, even if the share of income saved at all ages was the same. Ameriks and Zeldes (2000) show how a given set of age-wealth profiles over time can be consistent with very different underlying patterns of asset accumulation over the life-cycle as a result of different combinations of time and cohort effects.

If cohort and time effects play an important role in determining the observed pattern of asset holdings at different ages, one would expect to see significant differences over time in the cross-sectional age-wealth profiles. Figures 4 and 5 present these profiles from the 1989, 1995, and 2001 SCF. Figure 4 reports mean net financial assets, while Figure 5 reports medians for each age group. The figures show a level difference in mean financial assets, with 2001 greater than either 1995 or 1989 at all ages. Gale and Pence (2004) investigate the source of the increase in age-specific wealth, and find that changing household characteristics, such as higher rates of college completion, can explain a substantial part of the disparity. The shape of the age-net financial assets profile is nevertheless reasonably stable across the three different survey years.

Cross-sectional wealth profiles like those in Table 3 were used by Yoo (1994b) and Bergantino (1998) to summarize the potential evolution of asset demand as the population age structure changes. Mankiw and Weil (1989) followed a similar procedure in their analysis of real estate demand. These procedures are only valid under the assumption that cohort effects are equal for all age groups.

Nevertheless, a procedure like this can provide a useful starting point for evaluating how changing demographic structure will affect asset demand. Poterba (2001) uses repeated cross-sections of the Survey of Consumer Finances (SCF) from 1983, 1986, 1989, 1992, and 1995 to estimate age profiles of asset ownership allowing for different lifetime asset levels for different birth cohorts. Poterba and Samwick (2001) present related analysis focusing on individual asset categories. The specification controls for cohort effects but assumes that there are no time effects on net worth or net financial asset holdings. The estimated coefficients from this calculation are show in Table 4. They show that allowing for cohort effects has a surprisingly small impact on the estimated age structure of asset holdings. The basic age-wealth patterns for both net financial assets and net worth are similar in Tables 3 and 4. Table 4 also presents tabulations for common stock holdings at different ages, since much of the concern about asset market meltdown focuses on the equity market.

Estimates of the age-specific demand for financial assets such as those in Table 4 can be used to construct a measure of projected asset holdings per capita in each year, based on the age-specific structure of asset demands in a given year. This measure is defined by $\sum \alpha_i * N_{it}$ where α_i denotes the age-specific asset holdings and N_{it} denotes the actual or projected number of individuals of age i in year t . The resulting “projected asset demand” data series is shown in Figure 6. The figure shows that projected asset demand rises modestly over the four decade period between 1980 and 2020. There is little decline in asset holdings, however, after 2020. Projected asset demand does not decline as the “asset market meltdown” hypothesis predicts, largely because the underlying age-specific patterns of asset holding do not show a sharp decline at older ages. The next section investigates the historical correlation between the projected asset demand series in Figure 6 and the level of asset values. If asset supply is not perfectly elastic, high values of projected asset demand should be associated with low required returns and high asset prices.

One important limitation of calculations like those underlying Table 6 is that they focus on financial assets held directly by households. They omit assets held through defined benefit pension plans, which, even with the declining relative importance of defined benefit plans, still represent a substantial

share of the financial market. Schieber and Shoven (1997) point out that the mechanical draw-down of defined benefit pension assets in the years after the Baby Boom cohort reaches retirement will put downward pressure on asset prices. In most cases, the value of the assets that are accumulated in defined benefit plans peaks at the date when an individual retires. As benefits are paid out, the actuarial present value of the remaining payouts declines, and the assets needed to provide these benefits decline. This implies that there is a substantial force of accumulation and then decumulation as a large birth cohort ages. Current concerns about the long-term viability of some private sector defined benefit pension plans, and the possibility that these plans may be managed by the federal government through the PBGC, could reduce the importance of this automatic decumulation channel. It is nevertheless important in future work to consider the impact of defined benefit pension plans in tightening the link between demographic structure and the demand for financial assets.

Although cross-sectional age-asset profiles have well-known limitations, they can provide a starting point for analyzing how changing age structure in the population may affect the composition of asset ownership. Table 5 presents tabulations based on the 1989 and the 2001 Surveys of Consumer Finances. There are two panels, one for net financial assets and another for net worth. The first row in each panel shows the fraction either net financial assets or net worth held by families headed by someone between the ages of 20-39, 40-64, and over 65. In 1989, for example, over-65 families held 35 percent of net financial assets. The second row shows the predicted 2001 holdings, by age group, assuming that the 1989 age-specific pattern was unchanged, but allowing the population mix to change. The share of net financial assets held by those over the age of 65 was predicted to drop from 35 to 32.4 percent. The actual 2001 holdings of the over-65 group were 31.1 percent of net financial assets, so the 1989 cross-section over-predicted the actual 2001 holdings. The 1989 data under-predict net worth holdings in 2001. This provides some insight on the relative movements of net real estate assets and net financial assets.

The last two rows in each panel show the projections for 2020 and 2040 based on Census Bureau population projections and the 2001 cross-sectional pattern. The projections show a substantial increase in the share of assets held by the over-65 groups in both categories. For net financial assets, holdings are

predicted to rise to 44.3 percent, from 31.1 percent, for the over-65 group by 2040. The increase for net worth is similar, from 29.4 percent in 2001 to 42.3 percent in 2040. These statistics offer insight on the changing demographics of the market for financial services, and they suggest that the elderly will become more important and that younger households, those headed by someone under the age of 39, will become a much less important part of the financial services market by 2040.

5. New Evidence on Population Age Structure and Asset Returns

This section presents empirical results on the relationship between asset returns, the level of asset prices, and various measures of demographic structure in the United States. The analysis focuses on annual real returns on three assets -- Treasury bills, long-term government bonds, and large corporate stocks. Stock returns are measured by the return on the S&P 500 index. The explanatory variables consist of several measures of population age structure, as well as other control variables that might affect the level of asset prices. There are four demographic variables: the share of the total population between ages 40 and 64, the share of the total population over age 65, the share of the adult population between the ages of 40 and 64, and the share of the adult population over the age of 65. Some regression specifications include both the share of the population in middle age, and the share over the age of 65. Real returns are computed by subtracting each year's inflation rate, computed as the year-end to year-end change in the Consumer Price Index, from the pretax nominal return on each asset. The analysis focuses on the period 1926-2003, for which Ibbotson Associates (2004) provides reliable and comparable data on returns. For each of the three asset classes, the subsequent tables explore the link between demography and asset returns for the full sample period as well as for the post-war (1947-2003) sample. Studying several different asset categories provides information on returns on both relatively low-volatility assets and more risky assets.

5.1 Evidence on Asset Returns

Table 6 presents regression coefficients from equations that relate each asset return measure to the demographic variables. There are six sub-panels in the table, two for each of the three asset classes.

In each case there is one sub-panel for the 1926-2003 sample and one for 1947-2003. The results do not suggest important correlations between asset returns and demographic structure. There is weak evidence that in the fixed income markets, and particularly the Treasury bill market, population age structure is correlated with asset returns. The variable measuring the fraction of the population between the ages of 40 and 64 displays a statistically significant coefficient in the Treasury bill regressions for the full sample period. This coefficient on this variable is statistically insignificantly different from zero in the post-war sample regressions, and none of the other demographic variables have a statistically significant effect on asset returns. This is the basis for the claim that the evidence is weak, at best.

For the full sample period, the coefficient of -1.39 from a regression of real Treasury bill returns on the population share aged 40-64 implies implausibly large changes in the real return on Treasury bills. Between 2000 and 2040, Table 1 showed that the population share aged 40-64 is projected to rise by 0.18 percentage points. Multiplying this amount by the coefficient -1.39 in the first column of Table 6 suggests a decline of 240 basis points in real Treasury bill yields between 2000 and 2040. Effects this large seem particularly unlikely because the population share aged 40-64 rose by more than five percentage points between 1975 and 2000, without any analogous movement in the real return on Treasury bills. The coefficient estimates for the full sample, -1.39 on real bill yields and -1.73 on real bond yields, imply that a demographic change of this magnitude would reduce real bill yields by 650 basis points, and real bond yields by 900 basis points. These effects are much larger than the modest predictions from the simulation models discussed above. This suggests that the coefficients on the demographic variables may be affected by omitted variable bias, with other factors that are slowly varying influences on asset prices reflected in the estimated effects. A key limitation of the empirical analysis in Table 6, and with virtually all of the previous empirical analysis on demography and asset returns, is that it does not embed the analysis in a broader model of equilibrium asset return determination. As such, the equations lack control variables that might reduce the omitted variable problem.

The results for bond returns are similar to those for Treasury-bill returns. For stocks, however, the results are somewhat different. None of the coefficient estimates are statistically significantly different from zero. Moreover, the coefficient estimates suggest that increasing the share of the total population or of the adult population in the 40-64 age category raises equity returns. This is inconsistent with the simulation models described above in which expansion of the number of middle-aged workers raised demand for assets and reduced the prospective rate of return.

One potential concern with time series regressions like those in Table 6 is that there may be unit roots in the explanatory variables, since they are slowly-varying demographic time series. Poterba (2001) uses the tests proposed by Engle and Granger (1987) to test for unit roots in the residuals from estimated equations similar to those reported in Table 6. The results reject the null hypothesis of a unit root, and therefore lead to some confidence with the current results.

Several conclusions emerge from the empirical results. First, the most substantial correlation between demographic factors and population age structure obtains for Treasury bills. One explanation for this may be that real returns on bills are less volatile than other return series, so it is less difficult to detect the impact of a slowly-changing variable such as the demographic share in different age groups on these returns than on other time series. Nevertheless, there is no evidence that the real returns on corporate stocks for the last seventy-five years have been correlated with population age structure. Second, the demographic effect appears to be much larger in the pre-war period than in the postwar period. Studying the impact of the postwar Baby Boom cohort on asset markets does not provide any strong evidence of a link between demography and returns, even in the Treasury bill market. Finally, many measures of population age structure exhibit very little correlation with asset returns, so one must be careful in interpreting a finding that some demographic variable is correlated with returns. There is a danger of data-mining, in that the few statistically significant relationships may become the starting point for future research studies. This may lead to a spurious appearance of broad support for the link risk that the few specifications that yield statistically significant relationships become the standard workhorses for academic research, which then exaggerates the effective degree of correlation.

5.2 Evidence on Asset Price Levels

In simple theoretical models such as those discussed in section two, when a large age cohort begins to purchase assets for retirement, it bids up the price of capital. This implies a positive association between the level of asset prices and the demographic demand variable. To test whether such a relationship emerges in the data, Table 7 presents regression models that relate a measure of the price level for common stocks, proxied by the price-dividend ratio for the S&P 500, to various demographic variables. The dependent variable is the price-to-dividend ratio on the S&P500 at the end of the year.

The results in Table 7 provide some evidence of a link between population age structure and the level of asset prices. The first panel shows results in which demographic variables are the only explanatory variables. In the first column, the population share between the ages of 40 and 64 has a positive correlation with the P/D ratio, and the estimated regression coefficient is statistically significantly different from zero. The population share over age 65, the regressor in the second column, also shows a positive coefficient, although the magnitude is about half that of the population share between the ages of 40 and 64. This pattern remains in the third column, when both demographic variables are included. In the last three columns of the first panel, the demographic variables are scaled relative to the adult population. The coefficients on the share of the adult population between the ages of 40 and 64 and on the share over 65 remain statistically significantly different from zero, but now the relative magnitude of the coefficients reverses. The effect of an older population is greater than that of a middle-aged population.

The second panel in Table 7 adds control variables for the rate of economic growth and the real interest rate into the basic specification for the price-dividend ratio. Adding these control variables does not change the basic conclusions. The economic growth rate is measured as a three-year moving average of the growth in real GDP, and the real interest rate is the difference between the nominal interest rate on twenty-year Treasury bonds and the three year moving average CPI inflation rate. The coefficient on the growth rate is positive, which is consistent with the view that this growth rate proxies for the future

growth of dividends, but it is statistically insignificantly different from zero. The real interest rate is also always positive; this is inconsistent with simple valuation theory.

The unit root problems alluded to above are potentially more serious for the equations in Table 7 than for those in Table 6, because the dependent variable in Table 6, the real return, is nearly white noise, while the dependent variable in Table 7, the price-dividend ratio, is highly persistent. To address the econometric difficulties that might arise from regressors and dependent variables that are close to a random walk, the third panel of Table 7 reports coefficient estimates from a regression model in which the change in the price-dividend ratio is regressed on the change in the demographic measures from the earlier panels. In this case, the coefficient estimates for the demographic variables are no longer statistically significantly different from zero, and in each case when a measure of the elderly population is included, the coefficient estimate is negative rather than positive in the level specification. These results raise important doubts about the robustness of the findings in the upper panels of Table 7.

The three lower panels of Table 7 present similar regression equations estimated for the 1947-2003 sample. The results are broadly similar to those for the full sample period, and the results are again dependent both on the choice of the denominator for normalization, the adult population or the total population, and on whether the equation is estimated in differences or in levels.

The results in Table 7 suggest two conclusions. First, it is possible to use regression analysis to estimate large effects of demographic factors on the level of asset prices. While the point estimates of these effects are implausibly large, more plausible values usually fall within the 95 percent confidence interval for the coefficients. Thus there is more support for a link between demography and asset markets when we study the level of asset prices than when we study returns. Second, however, the results of this price level analysis may be subject to "spurious regression bias" because the dependent and explanatory variables are all slowly-trending time series. The coefficients from these models are sensitive to differencing, and to altering the sample period of estimation. The statistical significance of the results from differenced models is much lower than that from models estimated in levels, thus casting doubt on the findings.

The regression equations in Tables 6 and 7 all focus on the relationship between demographic variables and asset returns. Yet the discussion of age-specific asset holdings in the previous section, and the associated discussion of projected asset demand, suggests the possibility of a more subtle test that uses data on age-specific asset holdings. When the demographic variables shown in Table 6 are replaced with the projected asset demand variable, there is never a statistically significant relationship between projected asset demand and returns on Treasury bills, long-term government bonds, or corporate stock. With respect to the level of asset prices, there is some evidence of a positive association.

Table 8 reports regression equations relating the price-dividend ratio to the level of projected asset demand. The full-sample findings that relate the level of the price-dividend ratio to the level of projected asset demand, as well as those from the post-war subsample, suggest that an increase in projected asset demand is associated with an increase in the price-dividend ratio. This finding is robust to the choice between projected asset demand measured using age-specific net worth, net financial assets, or common stocks. As in some of the earlier specifications, the results again appear quite sensitive to the choice between the full sample and the most recent sub-sample, with the coefficient estimates doubling or tripling as a result of this sample change.

Table 8 also shows the effect of estimating differenced specifications for the projected asset demand equations, and once again the instability found in the earlier specifications emerges. None of the coefficient estimates from the differenced models are statistically significantly different from zero, although the point estimates continue to suggest a positive effect of projected asset demand on the price-dividend ratio. As with some of the earlier estimates, however, the coefficients imply larger demographic effects than theoretical analyses of demography and asset prices suggest. Consider, for example, the estimates of the link between projected net worth and the price-dividend ratio for the full sample period. The estimated coefficient is 1.89. The projected net worth variable equals 85.7 in 1980, and 92.9 in 2000. The coefficient estimate would therefore predict an increase of 13.6 in the price-dividend ratio between 1980 and 2000, solely as a result of demography-related changes in asset demand. The same equation predicts a further increase of 18 in the price-dividend ratio between 2000 and 2020.

The results in Table 8 raise an important question about why projected asset demand measures are more strongly correlated with the price-to-dividend ratio than simple measures of demographic structure. A key difference between the projected asset demand variables that constitute the explanatory variables in Table 8 and the simpler measures of demographic structure that were in earlier tables, is that the projected asset demand variables place roughly equal weight on retired individuals and prime-age workers. This is because the age-wealth profiles do not show substantial decline in old age. Thus the variables that seem to track at least the level of equity prices do not distinguish between prime-age workers and older individuals. This may be why Geanakoplos, Magill, and Quinzii (2004) find that the ratio of middle-aged individuals to young individuals, their “MY ratio,” has much greater explanatory power for the level of asset values than the other demographic measures that they consider.

6. Population Age Structure and the Composition of Financial Assets

The analysis in the last four sections has emphasized the potential effects of population aging on the demand for financial assets in aggregate, and the associated effects on financial market returns. A distinct question concerns the impact of these shifts on the types of financial assets, and financial services more generally, that households will demand. The aging of populations will shift emphasis from the accumulation of financial assets to the preservation of wealth and, for some households, the decumulation of wealth and the provision of insurance against late-life financial risks. This section presents a brief introduction to differences in age-specific probabilities of owning various asset categories, and then addresses changes that may be associated with population aging.

A natural starting point for analyzing how population aging will affect the demand for various assets is the current age-ownership profile for these assets. Table 9 shows the percentage of households headed by individuals of various ages who have various types of financial assets. The table shows that in 2001, the probability of stock ownership, either directly or through a retirement plan, was greater than fifty percent for households headed by someone between the ages of 30 and 59, and that it declined slightly for those in the early sixties and then by a somewhat greater amount at older ages. A smaller

share of households at all ages own bonds than own corporate stock. For bonds, however, there is again evidence of a decline in ownership probabilities at older ages.

It is tempting to extrapolate the age-ownership profiles in Table 9 and to draw the conclusion that an aging population will find fewer households holding stocks or bonds, either directly or through mutual funds. Yet this conclusion is dangerous, because there are difficult problems associated with age, time, and cohort effects, as described above. In the case of stock ownership, for example, there are good reasons to expect that the age-specific ownership rate at older ages will rise over time. As more households reach retirement age with assets in defined contribution pension plans such as 401(k) plans, there will be more elderly households with self-directed assets. Thus it seems likely that future equity and bond ownership rates at older ages will exceed the current rates. It is possible that the current and growing popularity of “lifecycle” funds in retirement accounts, which move the participant’s assets from equities to bonds as the participant ages, may lead to some automatic portfolio reallocation for retirement plan investors.

The problem of time-varying ownership probabilities can also be illustrated by reference to the home mortgage market. The third column of Table 9 shows the probability that a family that owns a home has a home mortgage. In 2001, there was a clear decline in mortgage probabilities from nearly 60 percent for those in the 55-59 age group to 30 percent for those in the 70-74 age group. This pattern, extrapolated forward, would suggest that as the population ages, there will be less demand for home mortgages. While this conclusion is likely to be broadly accurate, the imprecision of any possible forecast is illustrated by Figure 7. It shows the age-specific probabilities of mortgage indebtedness in the 1989 and 2001 SCFs. The figure shows that the two curves track each other closely at younger ages and through middle age, but then diverge around age 60. The probability of holding a mortgage was substantially higher for those in their sixties in 2001 than it was in 1989, presumably as a consequence of higher rates of refinancing during the 1990s than the 1980s, and the associated elongation of the number of years that households faced mortgage liabilities.

The problem of time-varying asset ownership probabilities is an important one and it needs to be recognized in any projection of future financial asset and liability holdings. It may nevertheless be of some interest to calibrate the changes in some aspects of financial asset demand that would be associated with changes in population age structure under the assumption that age-specific ownership patterns from 2001 persist into the future. Table 10 presents information on the fraction of owners of various assets who will fall into different age categories in 2020, and 2040, under this assumption. The table shows that in 2001, families headed by someone over the age of 65 accounted for 20.4 percent of all corporate stock investors. Table 11 shows that this group accounted for 34.8 percent of all corporate stock holdings. Table 10 then uses population projections to evaluate how this pattern will change prospectively. In 2040, the 2001 pattern suggests that 31.4 percent of all stockholders will be in families headed by someone over 65. This group is projected to hold 48.5 percent of all corporate stock. Thus, almost half of all corporate stock in 2040 would be projected to be held by families over the age of 65, up from roughly one third today. The ownership pattern for bonds, presented in the second panel of Tables 10 and 11, is similar.

The third panel of Tables 10 and 11 focuses on changing patterns of mortgage demand. Even though there are substantial demographic shifts, the aggregate ownership patterns for mortgages do not shift as much as the patterns for bond and stock ownership. The last sub-panel considers the changing pattern of demand for annuity products. Most of the demand for annuities occurs at older ages, so the shifting age distribution does increase the share of annuitants, and the fraction of annuities, held by those over the age of 65. The bottom row of Table 11 shows that in 2040, families headed by someone over the age of 65 are projected to hold 64 percent of all annuity contracts, up from fifty percent in 2001.

Annuities provide a useful illustration of how the aging population will affect the demand for financial services. The last column of Table 9 shows the percentage of households with an annuity contract. This percentage rises with age, from only two percent for households in their 40s to nearly ten percent for those in their 70s. As the population ages, it seems very likely that there will be increased demand for products like annuities that provide an income stream in old age and that also offer some

insurance against the financial risks that arise at advanced ages. In the case of annuities, they offer insurance against outliving one's financial resources. Mitchell et al. (1999) and Mitchell (2001) discuss the role of annuity markets in providing insurance, and the potential for selection effects to influence the attractiveness of this market for prospective buyers.

Long-term care insurance is another example of a product that provides insurance against late-life financial risks. In the retirement saving market, attention will shift from the focus on how to accumulate balances for retirement, to concern about how to move these balances to annuities, structured payout programs, and other withdrawal arrangements. The U.S. policy debate about retirement saving policy, which has been concerned largely with limits on retirement plan contributions, may begin to focus on minimum distribution requirements and related issues concerning payouts.

Annuities offer one mechanism for households that have accumulated financial assets to draw down their resources, but there are other mechanisms that are also likely to witness increased demand. Reverse annuity mortgages, for example, may become more popular as households seek to draw down their housing equity without moving. These products may have particular appeal for households that are cash poor but house rich, as many retirees in the highly-appreciated housing markets of the two coasts may be. Another device for translating asset wealth into cash payouts is cash dividend payments. Just as Goyal (2004) documents changing patterns of net cash inflows and outflows from the corporate sector as a function of population age structure, there may be growing demand for cash dividends as the population ages. Households with corporate stock portfolios also have other devices for translating their portfolios into cash, notably partial liquidation of their stock holdings. There may be growing demand for products that combine cash payouts with insurance of various types, such as long-term care insurance products that are paired with annuities, thereby providing life income if the buyer remains healthy and a nursing home benefit if the buyer is not well.

7. Conclusion

The correlation between asset returns on stocks, bonds, or bills, and the age structure of the U.S. population over the last seventy years, is weak. A statistically significant correlation emerges for Treasury bill returns in one sample period, but not another. The results are more favorable to the demographic hypothesis when the level of asset prices, as measured by the price-dividend ratio for the S&P500, is the dependent variable. The evidence is strongest when age-specific asset demands are used to construct time-varying "projected asset demands" at different dates, and when these "projected asset demands" are then related to price-dividend ratios, but even in this case the results are sensitive to changes in specification. None of the empirical findings provide a strong and convincing measure of the amount by which asset prices will change as the population of the United States and other developed nations ages.

The weak empirical findings stand in contrast to the results of most theoretical models that consider how demographic shifts will affect asset prices and asset returns. In most models, there are clear effects in predictable directions, with a baby boom driving up asset values and driving down returns for those in the large birth cohort. The simulation evidence from these models usually suggests a modest impact on asset values and returns, but it is nevertheless important to recognize that there is a strong theoretical presumption for these effects. Given the limited amount of time series data on returns and demographic variation, and the difficulty of controlling for all of the other factors that may affect asset values and asset returns, the theoretical models should be accorded substantial weight in evaluating the potential impact of demographic shifts.

A critical issue for policy-makers concerns with demographic change and asset markets is how policy variables may be brought to bear to offset any impact of changing population age structure. For example, if the monetary authority can affect the real interest rate on Treasury bills and long-term government bonds through its policy actions, then postulating a link between population age structure and equilibrium returns must make an implicit assumption about how the monetary authority would respond to changing age structure. The reaction of fiscal policy to population aging attracts much greater attention than

the reaction of monetary policy, largely because the fiscal policy effects of population aging have been discussed more widely than the consequences for asset markets.

Many studies have documented the fundamental impact that an aging population will have, in most developed nations, on the government's share of GDP and the fiscal deficit. Kotlikoff and Burns (2004) offer an overview of the issues facing the United States. The World Economic Forum Pension Initiative (2004) offers an introduction to related issues in other nations. The impact of demographic change on long-run fiscal balance depends critically on the future path of government transfer policy. In many nations, current policies are not sustainable, so some reform is inevitable. Will current commitments to deliver retiree health insurance be honored, or will they be modified? Will public pension programs, in the United States and in Europe, be modified so that future benefit obligations are more consistent with tax inflows? These questions of political economy are central to understanding how the demographic transition will affect population age structure, and in turn how the changing pattern of fiscal balances will affect financial markets.

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Table 1: Historical and Forecast Values for Indicators of Demographic Structure, 1920-2050

Year	Median Age	Average Age of Those 20+	Percent of Population 40-64	Percent of Population 65 and Over	(Population 40-64)/ Population 20+	(Population > 65)/ Population 20+
1920	25.3	40.3	22.2	4.63	0.375	0.078
1930	26.5	41.2	24.1	5.45	0.392	0.089
1940	29.1	42.2	26.5	6.84	0.404	0.104
1950	30.2	43.5	27.0	8.14	0.409	0.123
1960	29.4	45.3	26.5	9.23	0.431	0.150
1970	27.9	45.2	26.3	9.81	0.423	0.157
1980	30.0	44.5	24.7	11.29	0.362	0.166
1990	32.8	45.3	25.7	12.52	0.361	0.176
2000	35.4	46.7	30.1	12.43	0.422	0.174
2010	37.4	48.3	33.1	13.03	0.455	0.178
2020	38.1	49.6	31.4	16.27	0.424	0.221
2030	38.9	50.9	28.7	19.65	0.389	0.266
2040	39.0	51.5	28.3	20.42	0.384	0.276
2050	38.8	52.5	28.3	20.65	0.383	0.279

Source: U.S. Census Bureau historical data and projections from CPS Reports P25-1130. Average age over 20 computed using the midpoint in 5-year age intervals as the average age for all persons in that interval, and assuming that the average age for persons 85 and older is 90.

Table 2 Dependency Ratios in Developed Nations, 2000 and 2030

	Aged Dependency Ratio: Population 65+/Population 20-64	
	2000	2030
United States	0.208	0.365
Canada	0.205	0.411
Germany	0.263	0.498
Italy	0.290	0.506
United Kingdom	0.267	0.440
Japan	0.276	0.560
	Total Dependency Ratio: (Population < 20 or > 65)/Population 20-64	
	2000	2030
United States	0.695	0.808
Canada	0.624	0.796
Germany	0.601	0.799
Italy	0.604	0.771
United Kingdom	0.693	0.809
Japan	0.607	0.864

Source: World Economic Forum Pension Readiness Initiative (2004).

Table 3: Age-Specific Asset Holdings, 2001 Survey of Consumer Finances

Age of Household Head	Net Financial Assets		Net Worth	
	Mean	Median	Mean	Median
20-24	26330 (22642)	-340	44075 (29812)	3300
25-29	11649 (5103)	50	52282 (10098)	11895
30-34	32806 (10968)	940	88514 (14177)	20500
35-39	46504 (9065)	6300	122712 (22512)	37000
40-44	75099 (12506)	13540	204488 (24905)	68711
45-49	99240 (17412)	14000	240273 (29736)	74301
50-54	181181 (33148)	30130	369670 (57750)	103700
55-59	210908 (31985)	33450	455729 (64088)	134130
60-64	207848 (35873)	24000	421902 (75681)	109700
65-69	156288 (60076)	28525	346338 (75828)	119790
70-74	205077 (52811)	32800	409932 (103782)	133840
75 & up	174308 (8237)	27835	310900 (68733)	114000
All Ages	110185 (8237)	9850	240755 (14285)	59635

Note: Net financial assets subtracts consumer and investment debt from gross financial assets. Net worth is the sum of net financial assets, the gross value of owner-occupied housing, and holdings of other assets such as investment real estate, less the value of housing mortgage debt. All entries are measured in 2001 dollars. Standard errors are shown below the means.

Table 4: Age-Specific Asset Demands Estimated Allowing for Age and Cohort Effects, Surveys of Consumer Finances, 1983-1995

Age of Individual	Common Stock Holdings	Net Financial Assets	Net Worth
15-19	\$ 0 (0)	\$ 2285 (2823)	\$ 11042 (5391)
20-24	470 (134)	2170 (2939)	13656 (6337)
25-29	1477 (214)	4477 (3010)	25471 (6848)
30-34	3391 (367)	9402 (3126)	37706 (6648)
35-39	5906 (908)	14325 (3352)	60758 (7166)
40-44	10795 (1175)	20236 (4789)	86808 (7939)
45-49	18631 (1996)	37122 (4668)	123683 (10136)
50-54	23913 (2805)	57396 (6634)	151981 (15641)
55-59	32515 (3882)	71884 (7505)	177522 (17133)
60-64	31004 (4857)	80931 (8757)	189134 (19670)
65-69	30822 (5791)	92262 (9901)	201509 (22973)
70-74	28219 (7186)	92366 (11707)	173796 (25961)
75+	24722 (7482)	92239 (12091)	144316 (27026)

Notes: Estimates are based on regression models that relate real holdings of various assets by age cohorts in different survey years to a set of cohort “intercepts” and indicator variables for various age groups. Standard errors are shown in parentheses. See Poterba (2001) for further discussion. These estimates are the basis for the “projected asset demand” variable analyzed in regression tables.

Table 5: Past and Predicted Shares of Assets Held by Households of Different Ages, Survey of Consumer Finances 1989 and 2001

	Age of Household Head		
	20-39	40-64	Over 65
Financial Net Assets			
1989 Actual	13.1 (29.9)	51.9 (101.0)	35.0 (82.0)
2001 Predicted from 1989	10.6 (25.5)	57.0 (104.8)	32.4 (76.1)
2001 Actual	10.9 (16.0)	58.1 (71.7)	31.1 (36.2)
2020 Predicted from 2001	9.0 (12.9)	54.2 (65.8)	36.9 (42.1)
2040 Predicted from 2001	8.2 (11.8)	47.5 (56.7)	44.3 (49.6)
Net Worth			
1989 Actual	15.2 (34.1)	55.2 (106.6)	29.6 (69.2)
2001 Predicted from 1989	12.2 (29.1)	60.4 (110.4)	27.4 (64.1)
2001 Actual	12.0 (17.3)	58.6 (72.2)	29.4 (34.1)
2020 Predicted from 2001	9.9 (14.1)	55.1 (66.6)	35.0 (39.9)
2040 Predicted from 2001	9.1 (12.8)	48.5 (57.7)	42.3 (47.2)

Source: Author's calculations using cross-sectional age/net worth coefficients for household data from 1989 and 2001 SCF, along with estimates of the probability of household headship conditional on an individual being of a given age. These conditional probabilities estimated from the SCF were combined with Census projections on the future distribution of population by age to project future asset holdings.

Table 6: Demographic Structure and Real Returns on Financial Assets

1926-2003 Sample, Real Returns on Treasury Bills						
Population Share 40-64	-1.392 (0.357)		-1.530 (0.362)			
Population Share 65+		0.136 (0.205)	0.318 (0.190)			
Population 40-64/ Population 20+				-0.398 (0.817)		-0.392 (0.196)
Population 65+/ Population 20+					0.106 (0.157)	0.017 (0.160)
Adj. R-Squared	0.163	-0.008	0.184	0.046	-0.008	0.033
1947-2003 Sample, Real Returns on Treasury Bills						
Population Share 40-64	-0.632 (0.410)		-0.311 (0.372)			
Population Share 65+		1.050 (0.249)	1.003 (0.256)			
Population 40-64/ Population 20+				-0.298 (0.155)		0.077 (0.158)
Population 65+/ Population 20+					1.053 (0.213)	1.120 (0.255)
Adj. R-Squared	0.025	0.240	0.236	0.049	0.306	0.296
1926-2003 Sample, Real Returns on Government Bonds						
Population Share 40-64	-1.847 (0.983)		-2.049 (1.011)			
Population Share 65+		0.221 (0.527)	0.466 (0.529)			
Population 40-64/ Population 20+				-1.165 (0.474)		-1.261 (0.495)
Population 65+/ Population 20+					0.002 (0.403)	-0.283 (0.404)
Adj. R-Squared	0.034	-0.011	0.030	0.065	-0.014	0.058
1947-2003 Sample, Real Returns on Government Bonds						
Population Share 40-64	-1.331 (1.388)		-0.585 (1.356)			
Population Share 65+		2.424 (0.903)	2.335 (0.933)			
Population 40-64/ Population 20+				-1.142 (0.510)		-0.697 (0.602)
Population 65+/ Population 20+					1.935 (0.820)	1.327 (0.971)
Adj. R-Squared	-0.002	0.105	0.091	0.070	0.079	0.085

1926-2003 Sample, Real Returns on Common Stocks						
Population Share 40-64	1.425 (1.928)		1.368 (1.993)			
Population Share 65+		0.294 (1.013)	0.131 (1.044)			
Population 40-64/ Population 20+				-0.113 (0.949)		-0.102 (0.994)
Population 65+/ Population 20+					0.058 (0.774)	0.035 (0.811)
Adj. R-Squared	-0.006	-0.013	-0.020	-0.014	-0.014	-0.028
1947-2003 Sample, Real Returns on Common Stocks						
Population Share 40-64	2.880 (2.100)		3.428 (2.146)			
Population Share 65+		1.197 (1.462)	1.716 (1.477)			
Population 40-64/ Population 20+				-0.090 (0.816)		0.155 (0.977)
Population 65+/ Population 20+					0.596 (1.314)	0.731 (1.577)
Adj. R-Squared	0.016	-0.006	0.023	-0.019	-0.015	-0.035

Note: Each equation presents the results of estimating an equation of the form

$$R_t = \alpha + \beta*(\text{DEMOGRAPHIC VARIABLE})_t + \varepsilon_t.$$

Some equations include two demographic variables. Standard errors are shown in parentheses. Equations are estimated using annual data for the sample period indicated.

Table 7: Demographic Structure and the Price-Dividend Ratio on Common Stocks

1926-2003 Sample, Level Estimates						
Population Share 40-64	667.54 (79.46)		536.81 (74.00)			
Population Share 65+		375.56 (60.18)	248.55 (49.63)			
Population 40-64/ Population 20+				146.12 (67.43)		220.02 (53.86)
Population 65+/ Population 20+					275.66 (47.96)	311.35 (44.53)
Adj. R-Squared	0.475	0.330	0.601	0.046	0.294	0.415
1926-2003 Sample, Level Estimates						
Population Share 40-64	772.38 (69.80)		652.80 (68.35)			
Population Share 65+		363.94 (60.61)	193.67 (44.42)			
Population 40-64/ Population 20+				190.34 (67.86)		256.79 (53.49)
Population 65+/ Population 20+					266.95 (48.14)	303.19 (42.92)
Growth Rate	28.84 (30.71)	29.23 (41.06)	23.63 (27.57)	44.48 (47.53)	30.86 (42.08)	30.93 (36.93)
Real Interest Rate	206.59 (36.95)	71.25 (48.34)	171.48 (34.10)	148.94 (57.34)	77.16 (49.46)	127.06 (44.64)
Adj. R-Squared	0.626	0.332	0.699	0.101	0.298	0.459
1926-2003 Sample, Estimated in Differences						
Population Share 40-64	326.29 (296.33)		62.91 (371.38)			
Population Share 65+		-1722.17 (1068.86)	-1583.53 (1351.85)			
Population 40-64/ Population 20+				256.33 (171.97)		246.70 (174.64)
Population 65+/ Population 20+					-384.83 (646.39)	-257.91 (648.34)
Adj. R-Squared	-0.017	0.002	-0.011	-0.003	-0.028	-0.015
1947-2003 Sample, Level Estimates						
Population Share 40-64	780.21 (91.39)		737.46 (61.69)			
Population Share 65+		554.87 (111.87)	495.21 (59.72)			
Population 40-64/ Population 20+				148.38 (72.42)		348.50 (53.88)
Population 65+/ Population 20+					489.56 (104.04)	741.53 (88.19)
Adj. R-Squared	0.558	0.293	0.800	0.053	0.271	0.578

1947-2003 Sample, Level Estimates						
Population Share 40-64	819.32 (70.33)		762.70 (60.57)			
Population Share 65+		601.59 (161.37)	396.23 (83.15)			
Population 40-64/ Population 20+				275.09 (69.72)		377.96 (55.44)
Population 65+/ Population 20+					533.53 (158.73)	776.97 (122.28)
Growth Rate	-52.19 (68.50)	92.93 (122.44)	53.25 (61.94)	-227.45 (120.03)	15.93 (119.44)	-163.19 (91.84)
Real Interest Rate	279.02 (48.65)	-56.75 (114.31)	78.88 (58.75)	394.33 (88.84)	-43.49 (118.88)	33.01 (88.30)
Adj. R-Squared	0.740	0.274	0.815	0.292	0.245	0.590
1947-2003 Sample, Differenced Estimates						
Population Share 40-64	497.38 (322.21)		118.95 (552.60)			
Population Share 65+		-2060.32 (1171.91)	-1707.15 (2022.40)			
Population 40-64/ Population 20+				302.53 (177.25)		289.64 (180.10)
Population 65+/ Population 20+					-504.76 (679.57)	-357.84 (675.99)
Adj. R-Squared	-0.004	0.008	-0.010	0.005	-0.038	-0.008

Note: Each equation presents the results of estimating an equation of the form

$$(P/D)_t = \alpha + \beta*(DEMOGRAPHIC VARIABLE)_t + Z_t + \varepsilon_t$$

where Z_t denotes control variables such as the real interest rate or the three-year average of the GDP growth rate. When the equation is estimated in differences, the specification still includes a constant term. Some equations include more than one demographic variable. Standard errors are shown in parentheses. Equations are estimated using annual data for the sample period indicated.

Table 8: Projected Asset Demand and Price-Dividend Ratios

Sample Period	Age-Specific Asset Holding Measure Used to Construct Projected Asset Demand		
	Net Financial Assets	Common Stock	Net Worth
1926-2003, Levels	3.252 (0.502)	10.265 (1.585)	1.894 (0.282)
1926-2003, Differences	3.542 (4.546)	9.965 (10.393)	1.704 (1.794)
1946-2003, Levels	10.920 (1.068)	29.067 (3.399)	4.934 (0.576)
1946-2003, Differences	4.743 (4.961)	13.138 (11.154)	2.274 (1.922)

Note: Each entry shows the coefficient estimate of β from the equation

$$(P/D)_t = \alpha + \beta * (\text{PROJECTED ASSET DEMAND})_t + \varepsilon_t$$

where PROJECTED ASSET DEMAND is defined by using the 1983-1995 age-specific asset demand weights and time-varying population weights. Equations are estimated using annual data and standard errors are shown in parentheses. The differenced specifications relate the change in the P/D ratio to the change in projected asset demand, again allowing for an intercept term in the specification.

Table 9: Age-Specific Probability of Asset or Liability Ownership, 2001 Survey of Consumer Finances

Age of Household Head	Common Stock and Stock Mutual Funds	Bonds and Bond Mutual Funds	Mortgage, Conditional on Owning Home	Owner-Occupied Home	Annuity
20-24	30.3%	11.0%	85.3%	14.9%	0.2%
25-29	47.8	18.0	82.7	41.3	1.1
30-34	56.6	24.7	94.5	54.9	1.8
35-39	52.7	23.0	88.7	61.6	1.1
40-44	58.4	36.1	84.9	72.9	2.1
45-49	54.7	26.0	79.6	75.4	1.8
50-54	55.5	32.3	74.2	77.3	6.3
55-59	57.1	23.5	58.9	83.3	8.4
60-64	49.8	19.8	49.1	83.2	10.0
65-69	35.6	10.8	42.7	80.1	7.4
70-74	32.3	7.4	30.4	85.6	9.5
75 & up	31.0	10.0	11.5	76.0	8.0

Note: Entries for common stock and stock mutual funds, and for bonds and bond mutual funds, include assets held through defined contribution pension accounts as well as in traditional taxable accounts.

Table 10: Age-Specific Ownership Probabilities for Assets and Liabilities

	Age of Household Head		
	20-39	40-64	65 +
Common Stock and Stock Mutual Funds			
2001 Actual	29.99 (41.31)	49.66 (59.83)	20.35 (23.07)
2020 Predicted from 2001	25.87 (35.09)	48.77 (57.59)	25.36 (28.09)
2040 Predicted from 2001	24.55 (32.84)	44.05 (51.07)	31.40 (34.07)
Bonds and Bond Mutual Funds			
2001 Actual	14.02 (19.65)	55.71 (67.73)	30.27 (34.61)
2020 Predicted from 2001	11.57 (16.01)	52.34 (62.57)	36.08 (40.47)
2040 Predicted from 2001	10.67 (14.59)	45.93 (54.04)	43.41 (47.83)
Mortgages			
2001 Actual	33.91 (47.77)	57.12 (70.88)	8.97 (10.55)
2020 Predicted from 2001	30.30 (42.12)	58.12 (70.84)	11.58 (13.35)
2040 Predicted from 2001	30.08 (41.18)	54.92 (65.65)	15.00 (16.93)
Annuities			
2001 Actual	11.31 (16.06)	51.78 (62.92)	36.91 (41.99)
2020 Predicted from 2001	9.15 (12.87)	47.71 (57.15)	43.14 (48.28)
2040 Predicted from 2001	8.26 (11.51)	40.96 (48.55)	50.78 (56.15)

Source: Author's calculations based on age-specific asset ownership probabilities estimated from the 1989 and the 2001 SCF. Projections are based on Census population data as well.

Table 11: Age-Specific Ownership Shares for Assets and Liabilities

	Age of Household Head		
	20-39	40-64	65 +
Common Stock and Stock Mutual Funds			
2001 Actual	10.50 (17.57)	54.66 (70.61)	34.83 (41.99)
2020 Predicted from 2001	8.54 (14.15)	50.58 (64.32)	40.89 (48.42)
2040 Predicted from 2001	7.75 (12.76)	43.75 (54.95)	48.49 (56.60)
Bonds and Bond Mutual Funds			
2001 Actual	6.36 (21.70)	58.89 (89.31)	34.75 (49.48)
2020 Predicted from 2001	5.14 (17.50)	54.25 (80.88)	40.61 (56.72)
2040 Predicted from 2001	4.68 (15.90)	47.04 (69.21)	48.28 (66.42)
Mortgages			
2001 Actual	34.87 (49.52)	59.10 (74.11)	6.04 (7.39)
2020 Predicted from 2001	31.45 (44.14)	60.68 (74.93)	7.87 (9.48)
2040 Predicted from 2001	31.61 (43.77)	58.07 (70.45)	10.32 (12.21)
Annuities			
2001 Actual	5.24 (10.85)	44.35 (57.05)	50.40 (58.81)
2020 Predicted from 2001	4.08 (8.41)	39.28 (50.05)	56.64 (65.31)
2040 Predicted from 2001	3.54 (7.28)	32.41 (41.16)	64.06 (73.60)

Source: Author's calculations based on age-specific asset ownership shares estimated from the 1989 and the 2001 SCF. Projections are based on Census population data as well.

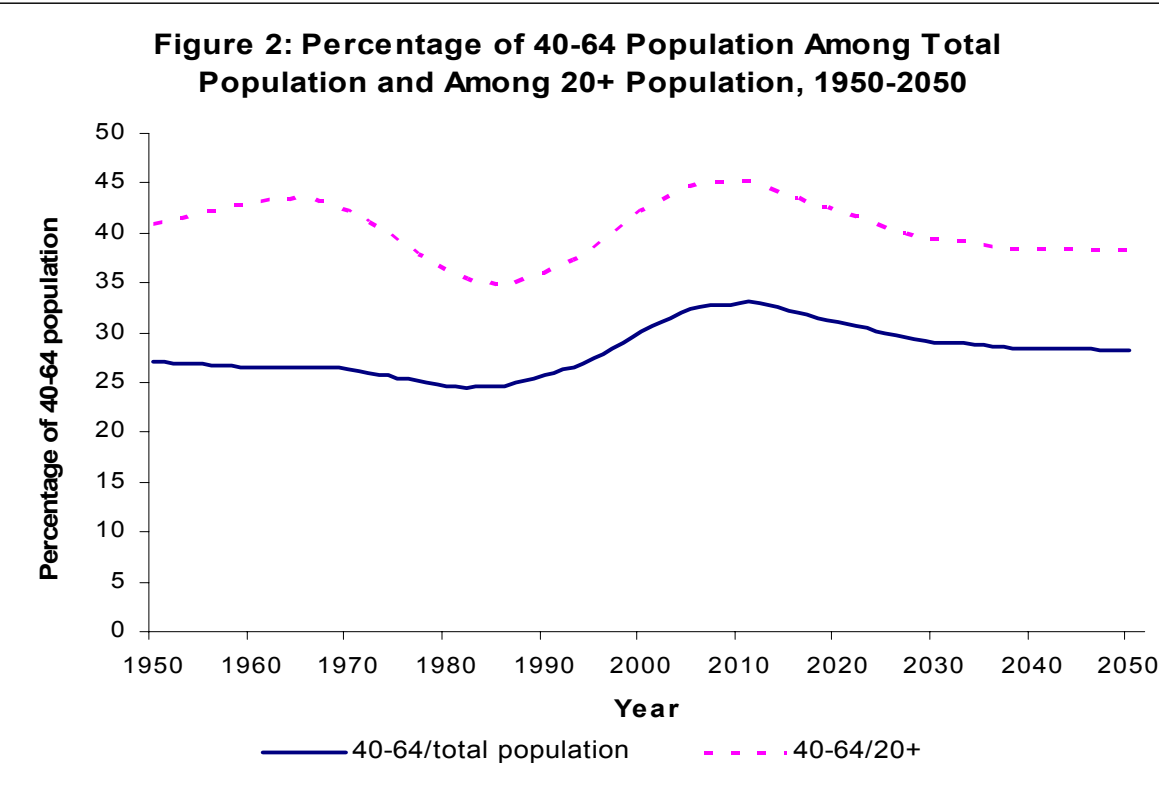
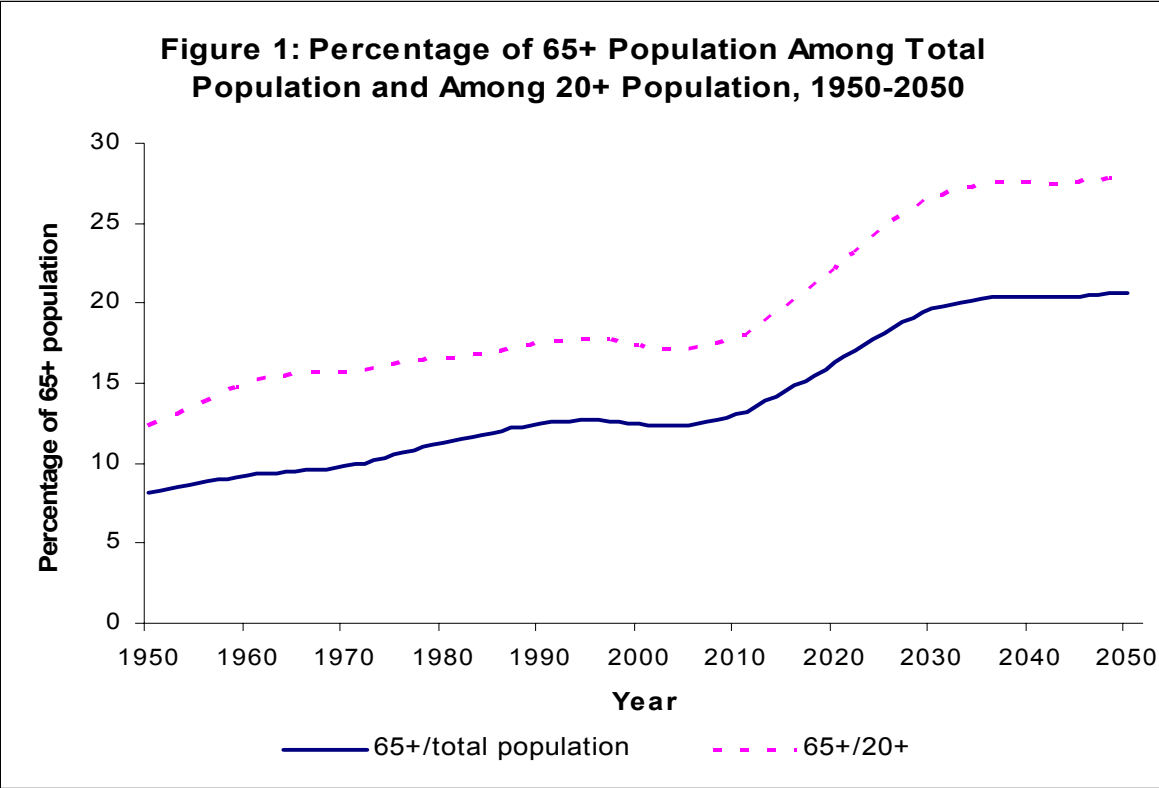


Figure 3: Real S&P500 Price Index and Percentage of 40-64 Population Among Total Population, 1950-2003

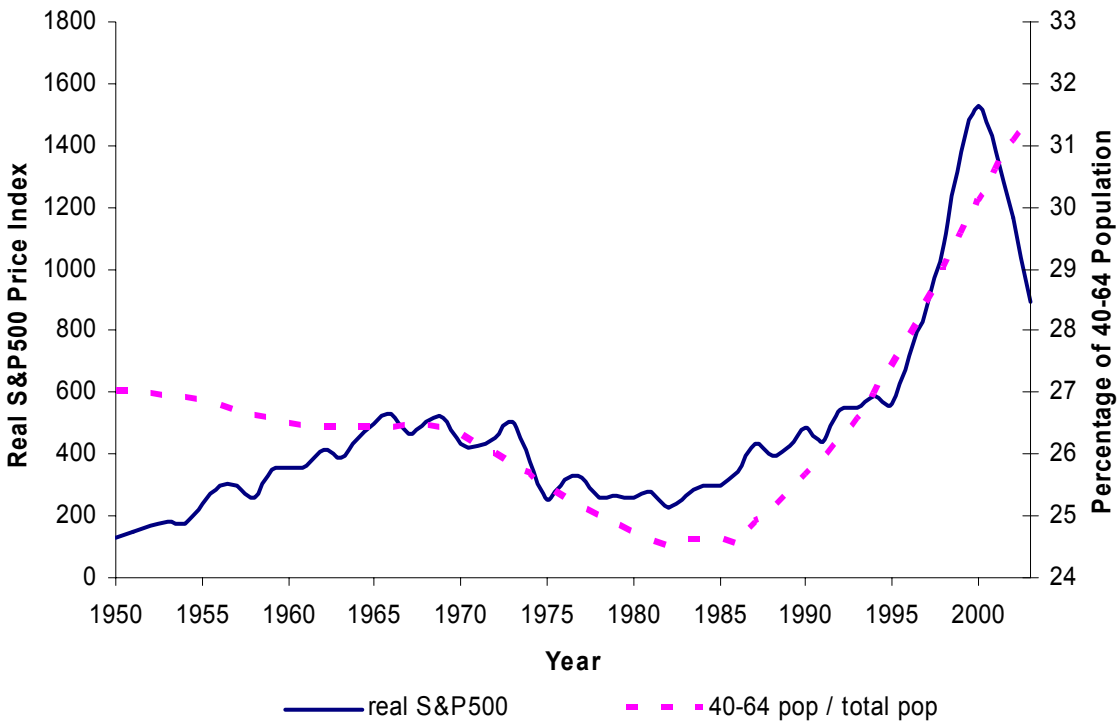


Figure 4: Mean Net Financial Assets by Age of Household Head, 1989-2001

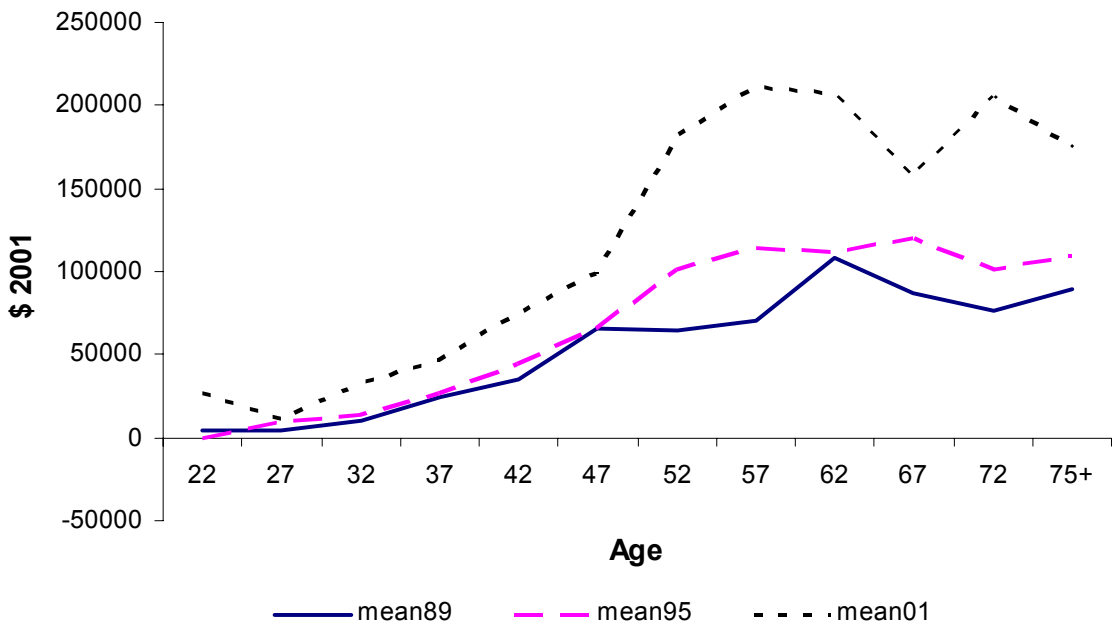


Figure 5: Median Net Financial Assets by Age of Household Head, 1989-2001

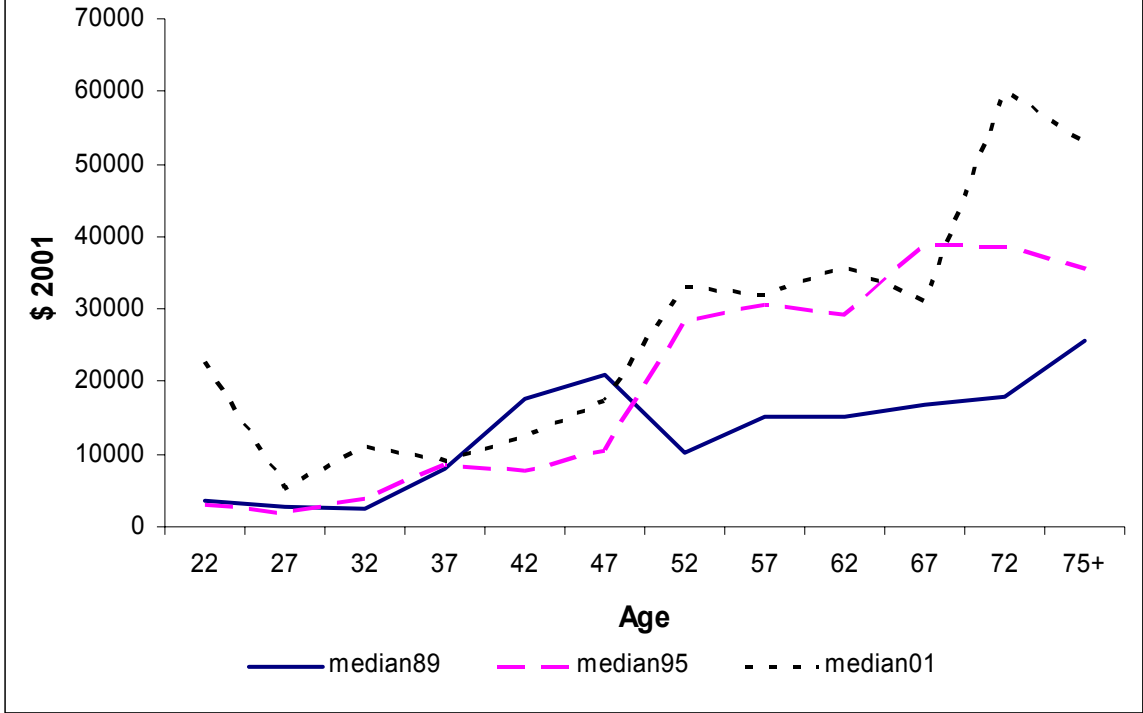


Figure 6: Projected Asset Demand Per Capita, Persons Aged 15 and Greater (\$2003)

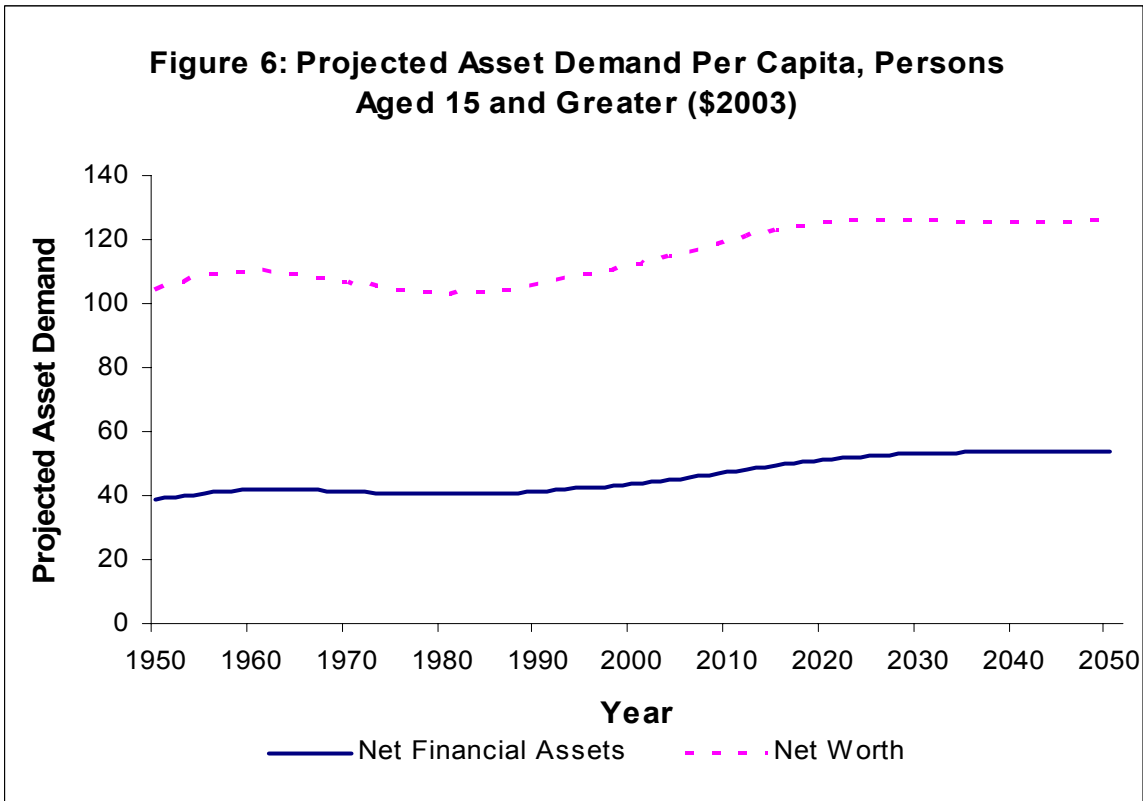


Figure 7: Probability of Mortgage Indebtedness for Homeowner Households of Different Ages, 1989 and 2001

